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Poaching of protected wolves fluctuated seasonally and with non-wolf hunting

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Poaching is the main cause of mortality for many large carnivores, and mitigating it is imperative for the persistence of their populations. For Wisconsin gray wolves (*Canis lupus*), periods of increased risk in overall mortality and poaching seem to overlap temporally with legal hunting seasons for other large mammals (hunting wolves was prohibited). We analyzed monitoring data from adult, collared wolves in Wisconsin, USA (1979–2012, n = 495) using a competing-risk approach to test explicitly if seasons during which it was legal to train hunting hounds (hounding) or hunt other large mammals (hunting) affected wolves' hazard of cause-specific mortality and disappearance. We found increases in hazard for disappearances and documented ('reported') poaching during seasons with hunting, hounding or snow cover relative to a season without these factors. The 'reported poached' hazard increased > 650% during seasons with hunting and snow cover, which may be due to a seasonal surge in numbers of potential poachers or to some poachers augmenting their activities. Snow cover was a major environmental factor contributing to poaching, presumably through increased detection of wolves. Our study suggests poaching is by far the highest mortality hazard for wolves and reinforces the need for protections and policies targeting poaching of protected populations.

Humans threaten the survival of large carnivores and the viability of their populations through habitat loss, killing and prey depletion¹. Consequently, the contraction, depletion and extirpation of large carnivores has contributed to simplification of trophic structures linked to both lower biodiversity and degraded ecosystem functions^{1–3}, suggesting the elimination of large carnivores “is one of the most significant anthropogenic impacts on nature”^{1,3}. Moreover, there is a growing concern for the wellbeing and claims of individual nonhuman animals and large carnivores within conservation^{4–6}. Increased consideration of nonhuman claims demands robust assessments of how anthropogenic activities, including those aimed at other species, impact risk of harm, including death^{7,8}.

Importantly, poaching, both reported and cryptic, is the main form of anthropogenic mortality for various regions' carnivores^{9–14}, including four US wolf (*Canis lupus*, *Canis rufus*, *Canis lupus baileyi*) populations^{15–18}. Here, we distinguish between these two poaching variants by their detection on the landscape, following^{12,15,17,18}; while 'reported poaching' refers to the component of total poaching that is reported, evidenced and thus detected by management agencies, 'cryptic poaching' refers to poaching that remains concealed and thus undetected. The concealment of poaching (its cryptic component) contributes to its systematic underestimation^{12,15,17–19}, increasing concerns over the viability of large carnivore populations subject to additional sources of anthropogenic mortality^{20–23}. Given both its prevalence and cryptic nature, mitigating poaching seems imperative for the persistence of many large carnivore populations, including endangered ones that are not subject to hunting seasons^{10,11,16–18,24}.

For wolf populations in the US, recent research has explored the effect of reducing protections for the species on cause-specific mortality, including poaching and its cryptic variant. Invariably, such studies have found an increase in poaching risk or incidence during policy time periods when species protections are reduced; i.e., when targeted lethal management by agency personnel, rather than unselective public hunting seasons, is sanctioned^{17,18}. For Wisconsin wolves, results are largely consistent with research detecting unmeasured mortality necessary to account for the slowdown in population growth during periods of reduced protections in that population^{25–27}. Relative to full protection periods, wolves in Minnesota also face an increased risk of overall anthropogenic mortality and poaching once protections are reduced, including public hunting, even if protections are later reinstated²⁸.

Research on intra-year mortality risk for Wisconsin wolves also found that periods of increased risk in overall mortality and poaching overlapped with hunting seasons for other large mammals, such as white-tailed deer

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Season	Hunt/hound			Hunt/hound/snow			Snow		
	HR (se)	95 CI		HR (se)	95 CI		HR (se)	95 CI	
LTF	1.18 (-0.30)	0.72	1.95	1.19 (-0.37)	0.65	2.19	1.52 (-0.40)	0.90	2.55
Legal	1.78 (-0.75)	0.79	4.05	0.72 (-0.69)	0.11	4.71	0.00 (.)	0.00	0.00
Reported poached	1.23 (-0.54)	0.52	2.91	7.58*** (-3.34)	3.19	17.99	3.27*** (-1.46)	1.36	7.86
Natural	238.98*** (-405.19)	8.61	6630.66	392.14*** (-709.72)	11.29	13,614.75	623.97*** (-1429.70)	7.00	55,655.22
Unknown	1.20 (-0.80)	0.33	4.44	7671.73*** (-24,384.75)	15.11	3,894,537.72	0.66 (-0.54)	0.13	3.32
Collision	2.61 (-2.24)	0.49	14.06	0.23 (-0.30)	0.02	3.00	1.17 (-1.09)	0.19	7.27
tv_c - (ln(t))									
Natural	0.39*** (-0.11)	0.23	0.66	0.44*** (-0.13)	0.25	0.77	0.41** (-0.14)	0.20	0.82
Unknown	-	-	-	0.17*** (-0.10)	0.05	0.55	-	-	-

Table 1. Hazard ratio (HR) point estimates from the stratified (by endpoint and protection period) joint Cox Model 3 (our best performing model, see Supplementary Material Tables 3–7 for model statistics, diagnostics and other models) for $n = 495$ monitored adult wolves, by endpoint and season (LTF = ‘lost to follow-up’, defined in Methods). We present HRs and compatibility intervals (95 CI) for all endpoint-season interactions relative to a baseline season. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

(*Odocoileus virginianus*) and black bear (*Ursus americanus*), and hypothesized such increases in poaching risk were in part attributable to the surge of hunters on the landscape during those periods^{14,18,29}. Similarly, in Minnesota, Nov–Apr is the period of highest overall anthropogenic and illegal killing of wolves, with the authors again pointing to the overlap with firearm season for white-tailed deer²⁸. Critically endangered red wolves in the Southeastern US also face increased risk of anthropogenic mortality (mostly attributable to gunshot) and disappearances during fall and winter hunting seasons for other large mammals^{16,24}.

The estimated increases in anthropogenic and illegal killing of wolves during other large mammal hunting seasons is also supported by social science research on hunter motivations and inclinations to poach wolves. Various surveys of Wisconsin residents spanning over a decade, and two qualitative focus groups, revealed rising inclinations to poach after federal protections were reduced and the state sanctioned lethal management^{30–32}. Treves et al.³⁰ found that increased inclination to poach wolves was correlated with perception of competition over deer, rather than fear or loss of domestic animals. Moreover, a quarter of bear hunters in that study said they would poach a wolf. Subsequent focus group research revealed that bear hunters generally hold negative attitudes towards wolves and wolf management, and that they “...believe that bear hunters, in general, sanction the illegal killing of wolves”³¹, p. 6. Farmers’ attitudes toward wolves did not differ significantly from those of hunters, and they believed that most farmers “approved, or were at least tolerant, of illegally killing wolves” (p. 6) The same study revealed deer hunters hold a range of attitudes towards wolves, significantly more positive than farmers or bear hunters, yet with some endorsement or participation in their illegal killing. Later survey research by Hogberg et al.³² highlighted a continuing negative trend in attitudes among male respondents and hunters living in wolf range before and after the state’s first legal hunt in 2012. All studies found net shifts towards agreement with the perception that wolves threaten deer hunting opportunities.

In this study, we analyze monitoring data from adult, collared wolves in Wisconsin, USA (1979–2012, $n = 495$ collared adults) to test explicitly if seasons during which it was legal to train hunting hounds (hounding) or hunt other large mammals (hunting wolves was prohibited; see “Methods” section) affected wolves’ hazard of cause-specific mortality and disappearance (endpoints hereafter). Our explicit modelling of intra-wolf-year anthropogenic and natural seasons allows us to explore any interactions between endpoints within seasons, as well as interactions between anthropogenic and natural landscape conditions (e.g., simultaneous hunting and snow cover). Our results suggest poaching hazard, both cryptic and reported, is substantially higher during seasons with hunting and snow cover relative to seasons without these factors. Our methods can promote the conservation and consideration of wild animals through improving the evaluation of anthropogenic impacts on their mortality and disappearances, as well as the effectiveness of policies aimed at protecting them and mitigating poaching.

Results

Estimating unconditional, endpoint-specific hazards. We built 3 stratified (by endpoint and protection period [lib_kill]), joint Cox models (see model statistics in Supplementary Material Table 3). We present results by endpoint for our best model (Table 1), following our model selection criteria (see Supplementary

Material Tables 4–6 for results from Models 1 and 2). Results, largely consistent across models (Table 1, Supplementary Material Tables 4–6; and see Supplementary Material Table 7 for analogous ‘known-LTF’ model), reveal that both anthropogenic and natural seasons were associated with meaningful increases in the hazard of multiple endpoints for collared adult wolves, especially of reported poached.

Lost-to-follow-up (LTF). Hounding and hunting seasons without snow (hunt/hound; Jul–Nov 14th) were associated with an 18% (HR 1.18, 95 CI 0.72–1.95) increase in hazard of LTF relative to the baseline period (April 15th–June). Similarly, the hunting and snow season (hunt/hound/snow) increased the hazard of LTF by 19% (HR 1.19, 95 CI 0.65–2.19). The snow season outside hunting or hounding periods (snow) increased the relative hazard of a wolf going LTF by 52% (HR 1.52, 95 CI 0.9–2.55).

Legal. Snowless hounding and hunting seasons (hunt/hound) were associated with a 78% increase in hazard of legal killing for wolves (HR 1.78, 95 CI 0.72–1.95), relative to the baseline season. On the other hand, hounding and hunting seasons with snow (hunt/hound/snow) decreased the hazard of a wolf being killed legally by 28% (HR 0.72, 95 CI 0.11–4.71). There were no records of wolves being killed legally during the snow season.

Reported poached. The hunt/hound season increased the hazard of wolves being reported poached by 23% (HR 1.23, 95 CI 0.79–4.05). The hunt/hound/snow period was associated with the highest hazard of wolves being reported poached, with a substantial increase of 658% over the baseline season (HR 7.58, 3.19–17.99). The snow season without hounding or hunting (snow) was associated with another substantial, albeit lower than with hunting, increase in hazard for wolves being poached and reported, this time by 227% (HR 3.27, 95 CI 1.36–7.86).

Natural, unknown and collision. The hazard of a natural endpoint showed substantial initial increases in hazard relative to baseline for all seasons, but with considerable non-proportional decreases in hazard with monitoring time (Table 1). The natural endpoint saw higher increases in hazard during the snow (HR 623.97, $\text{tvc}=0.41$) and hunt/hound/snow (HR 392.14, $\text{tvc}=0.44$) seasons than for the snowless hunt/hound season (HR 238.98, $\text{tvc}=0.39$). The hazard of an unknown endpoint increased during hunt/hound (HR 1.2), increased during hunt/hound/snow seasons but with a considerable non-proportional decrease over time (HR 7671.73, $\text{tvc}=0.17$), and decreased during the snow season (HR 0.66). The hazard of wolves dying by collisions increased during the hunt/hound (HR 2.61) and snow seasons (HR 1.17), and decreased during the hunt/hound/season (HR 0.23). The low number of events per variable (EPV, see Methods) for both the unknown and collision endpoints reduce our confidence in their results.

Analysis of cumulative hazards and incidences over monitoring time ‘t’, by season. Below we present results of constructed cumulative hazard curves (Fig. 1, Panels A–C) and CIFs (Fig. 2, Panels A–C) using our stratified joint Cox Model 3 (Table 1). Figure 1 illustrates how endpoint-specific hazards accumulate over a wolf’s monitoring time, by season, and allow for comparing the magnitude (rather than HR) of each endpoint’s hazards. Figure 2 allows for discerning any interactions between endpoint hazards over time.

Baseline season (Figs. 1, 2, panels A–C). LTF has by far the highest cumulative hazard and incidence of all endpoints throughout the season. Both hazard and incidence of other endpoints are much lower relative to LTF. The second highest cumulative hazard during the season belongs to reported poached until $t=600$ (Fig. 1), when it is matched by the hazard of a natural endpoint (lower before) up to $t=1200$ (Fig. 1). The hazard of a natural endpoint becomes the second highest cumulative hazard during the season at $t>1200$ (Fig. 1), yet its incidence remains lower than reported poached until $t=2000$, when it reaches similar levels (0.12, Fig. 2).

Hunt/hound season (Figs. 1, 2, panel A). LTF remains the endpoint with the highest cumulative hazard and incidence of all endpoints (Figs. 1, 2) throughout the hunt/hound season, despite having the lowest HR increase (Table 1). The legal killing (during strict protection periods) and reported poached endpoints (both with $\text{HR}>1$, Table 1) have the second largest, and similar, cumulative hazards up to $t=700$ (Fig. 1), after which reported poached overtakes legal killing as the second largest cumulative hazard (despite the lower HR). However, both endpoints maintain similar levels of incidence throughout t . The increase in hazard of legal killing results in an increased incidence ($0.12-0.085=0.35$, $t=2000$; Fig. 2) similar in magnitude to the observed decrease in cumulative incidence of LTF ($0.562-0.525=0.037$, $t=2000$; Fig. 2), which suggests the decrease in LTF incidence ($\text{HR}>1$, Table 1) is influenced by the increase in hazard and incidence of legal killing. This increase in legal killing hazard may also preclude higher increases of incidence of the reported poached endpoint, despite the latter also having an $\text{HR}>1$ (Table 1). The cumulative hazard of a natural endpoint becomes lower than during the baseline season by $t<450$ (Fig. 1), and is the lowest cumulative hazard in the season throughout t . The incidence of a natural endpoint equals that of legal killing and reported poached until $t=700$ (Fig. 2), after which it becomes the lowest.

Hunt/hound/snow season (Figs. 1, 2, panel B). Reported poached is the endpoint with the highest cumulative hazard throughout t , followed by LTF (Fig. 1). Indeed, the reported poached cumulative hazard is more than 1.5 times the cumulative hazard of LTF by $t=750$ ($0.81/0.53=1.52$) and until $t>1200$ (Fig. 1). Reported poaching also has the highest incidence throughout t . Figure 2 shows how the magnitude of the reported poached hazard results in a substantial increase in incidence of the endpoint, but also suggests the reported poaching hazard may

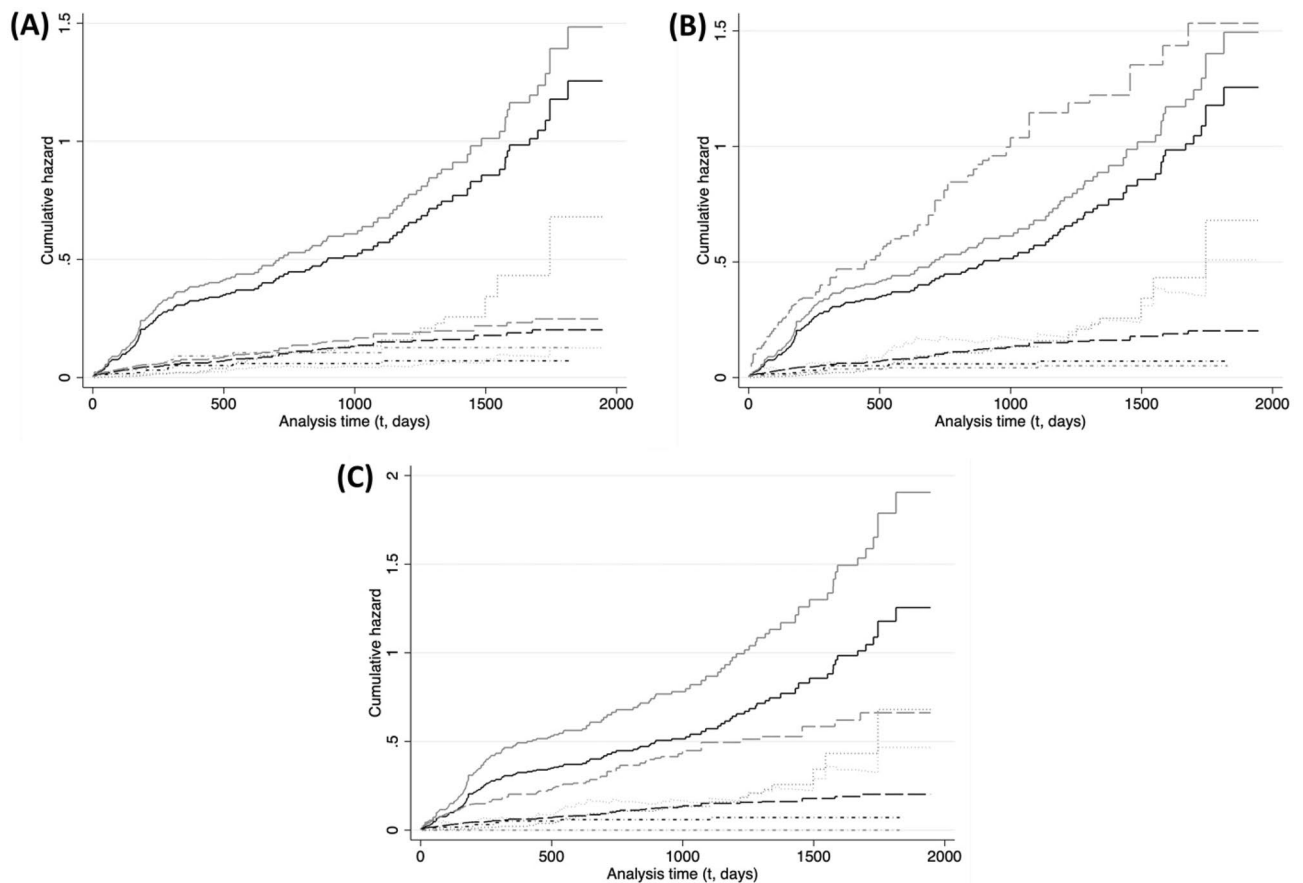


Figure 1. Endpoint-specific cumulative hazard over monitoring time (in days) during strict protection periods ($lib_kill=0$) derived from endpoint-season specific hazards obtained from our preferred joint stratified Cox model (Model 3, Table 1) for $n=495$ adult monitored wolves in Wisconsin, USA (1979–2012). Each panel corresponds to a season ((A) *hunt/hound*, (B) *hunt/hound/snow*, (C) *snow*) and illustrates the baseline (black curves) and seasonal (gray curves) cumulative hazards for our endpoints of interest: LTF (solid), reported poached (longdash), legal killing (dash-dot) and natural (dot).

play a role in the observed decrease in incidence of LTF (despite an $HR > 1$, Table 1). The third highest cumulative hazard and incidence throughout the season belongs to the natural endpoint, which is only on average about a third of the LTF cumulative hazard (Fig. 1) and half its incidence (Fig. 2). Legal killing has the lowest cumulative hazard throughout t .

Snow season (Figs. 1, 2, panel C). LTF has the highest cumulative hazard and incidence throughout t . The second highest cumulative hazard belong to the reported poached endpoint, which amounts on average to half of the LTF cumulative hazard throughout t . The natural endpoint has the third highest cumulative hazard in the season but the second highest incidence, which is marginally higher than reported poached throughout t (Fig. 2). This increase in natural incidence relative to baseline, despite similar cumulative hazards, may be due in part to decreases in hazard of the unknown and legal endpoints (Table 1).

Discussion

Time-to-event models for wild animals generally model exposure of individuals to natural conditions that may affect the risk of mortality and disappearance. Most models neglect to consider seasons of high human activity that may affect such risks, or interactions between endpoint hazards (reflected in incidences) that may illuminate ecology. For many large carnivores, which suffer from low natural mortality yet are also subject to high risk of anthropogenic mortality and poaching, seasons of anthropogenic activity may be as important as natural ones in mediating cause-specific mortality and disappearance.

Importantly, such anthropogenic seasons of higher mortality need not be specific to the animals being studied, especially if the species is controversial and much mortality illegal: our anthropogenic seasons consist of state hunting and hounding seasons for species other than wolves (i.e., deer or bear hunting, and hounding; not wolf hunting), but that mediate human activity on the landscape during those seasons. Our results support the hypothesis that increases in poaching risk during hunting seasons may be attributable to the surge of individuals with inclination to poach on the landscape^{14,18,29}. Alternatively, it could also suggest enhanced criminal activity of a few poachers during the same periods. We temper this increase in poaching risk by establishing snow cover as a

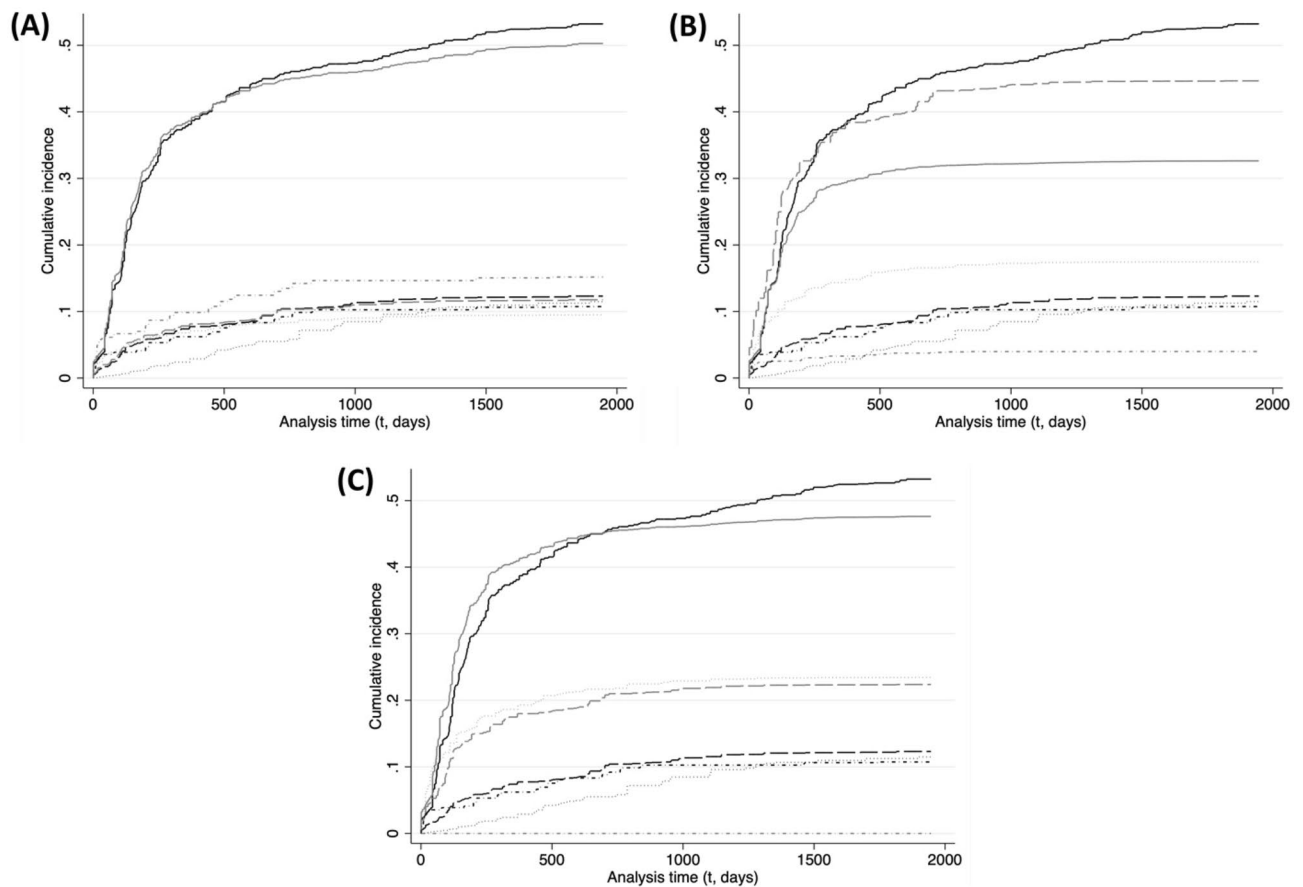


Figure 2. Endpoint-specific cumulative incidence curves (CIFs) over monitoring time (in days) constructing using all endpoint hazards obtained from our preferred joint stratified Cox model (Model 3, Table 1) for $n = 495$ adult monitored wolves in Wisconsin, USA (1979–2012). Each panel corresponds to a season ((A) *hunt/hound*, (B) *hunt/hound/snow*, (C) *snow*) and illustrates the baseline (black curves) and seasonal (gray curves) cumulative incidences for our endpoints of interest: LTF (solid), reported poached (longdash), legal killing (dash-dot) and natural (dot).

major environmental factor strongly associated with poaching. Moreover, our time-to-event analyses illuminate how to evaluate the effects that such anthropogenic seasons may have on risk of mortality and disappearance of monitored animals throughout their lifetime, and how considering such seasons may elucidate the mechanisms behind anthropogenic mortality and disappearance.

Additionally, our analysis period precedes and completely excludes any established public wolf hunting seasons. Hence, our modeled anthropogenic seasons represent the periods of most relevant anthropogenic activity for wolves, as hypothesized by other studies^{14,29,33} and suggested by social science studies on inclinations to poach self-reported by both deer hunters and bear hunters, as well as acceptance of poaching by hunters and farmers^{30–32}.

Our analyses show increases in the hazard of disappearances of collared wolves (LTF) relative to the baseline period (which excludes environmental and anthropogenic risks) for all seasons. The highest hazard of LTF occurs during the snow season, whereas increases in hazard are lower (and similar) for the two seasons that included hounding and hunting. LTF may experience changes in hazard due to changes in the hazard of any/all of its components: migration, collar failure, or cryptic poaching.

Constant and steep increases in LTF hazard throughout a wolf's lifetime suggests mechanisms other than migration regulating LTF hazard, given migration for adults is most frequent by yearlings and younger adults, around 1.5 to 2.2 years^{34–36}. Moreover, only migration out of state would end monitoring, not routine extraterritorial movements of radio-collared wolves. That our seasonal LTF curves depict the cumulative hazards more than doubling beyond those t generally associated with dispersal ($\sim t < 500$, given wolves were collared as adults), and that such hazards remain high throughout a wolf's lifetime relative to other endpoints, suggests mechanisms behind LTF hazard that are additional to migration out of state. If migration had been the driving mechanism behind LTF hazard, we would also expect higher increases in hazard (more similar to the snow season) during other periods also associated with increased dispersal for adults, such as Oct–Nov³⁶ within the hunt/hound and hunt/hound/snow seasons. Instead, during the latter seasons we observe smaller increases in LTF hazards, again suggesting mechanisms other than long-range movements out of state raising LTF hazard.

Although our study is unable to evaluate the contribution of collar failure to LTF hazard, we note that average and max time to LTF ($t = 497, 2330$ respectively) was similar to that of other anthropogenic endpoints (legal, $t = 472, 2357$; poached, $t = 477, 2303$) and much shorter than for other endpoints (collision, $t = 590, 2235$; natural, $t = 655, 3051$; unknown, $t = 773, 2999$) or censored observations ($t = 882, 2833$), which implicates causes other than battery or collar failure¹⁴.

As for the cryptic poaching component of the LTF hazard, the mechanism is consistent with the observed steep increase in hazard of LTF throughout a wolf's lifetime and seasons (contrary to the natural hazard), and with the similarities in time to endpoint between LTF and other anthropogenic, intentional killing (i.e., legal killing and reported poached).

The lower increases in LTF hazard during the hunt/hound and hunt/hound/snow seasons relative to the snow season show different patterns to that of the reported poached endpoint. We hypothesize the much higher relative cumulative hazard of the LTF endpoint for all seasons except hunt/hound/snow (for which reported poached is highest) may suggest a rate of cryptic poaching that increases not only due to more cryptic poaching activity than baseline during periods of more anthropogenic activity (hunt/hound and hunt/hound/snow seasons), but also due to decreased detection of poaching on the landscape given environmental conditions during the snow season³³. This reduced detection of cryptic poaching which increases LTF hazard during the snow season does not translate to the hunt/hound/snow season (despite similar environmental conditions) due to a surge of individuals on the landscape that result in not only more, but detectable poaching, therefore increasing the reported poached rather than the LTF hazard. This seems to resemble the pattern reported in Santiago-Ávila et al.¹⁸ of an increase in the hazard of reported poached relative to that of LTF during a census period in which dozens of civilian wolf-trackers went out in snow months to count wolves. Therefore, search effort and visibility due to landscape conditions are important variables to consider when designing anti-poaching interventions.

The hazard of reported poached more than doubles during the snow season relative to the baseline season, and doubles again during the hunt/hound/snow season, during which wolves are simultaneously exposed to environmental and anthropogenic conditions. The reported poached cumulative hazard during the hunt/hound/snow season is the highest of any across endpoint-seasons. These results implicate snow cover as a major factor mediating poaching activity (much lower hazard during snowless seasons), potentially by increasing wolf track detection. To those conditions, the hunt/hound/snow season may add more potential poachers or their increased killing, particularly during the (firearm) deer season, which more than doubles the snow season reported poached hazard. An important observation is that despite a decrease in incidence of LTF that season, in fact the LTF hazard increases, which points to this seasonal decrease in LTF incidence being an effect of the substantial increase in reported poaching hazard; i.e., the much higher rate of reported poached decreases LTF incidence despite an increased hazard of LTF. Therefore, we conclude that the reporting and documentation of poaching is improved when there are more people on the landscape, and worsened when there are fewer and snow cover is high.

For all anthropogenic and environmental seasons modelled, the natural endpoint shows an initial higher hazard but with a decrease in its seasonal hazard over time relative to baseline (i.e., non-proportional effects). The natural hazard is in general lowest during the hunt/hound/snow and snow seasons, the natural hazard is substantially lower than the LTF or reported poached endpoints. Moreover, the deceleration in the increase in natural hazards relative to the baseline period is suggestive of wolves learning to mitigate some seasonal natural hazards over their lifetime (e.g., intraspecific strife, starvation). We do not observe such a pattern with the LTF or reported poached endpoints, for which increases in hazard continue unabated over time. The difference in patterns between natural and anthropogenic endpoints suggests wolves may have difficulty and limited success in mitigating the hazard of anthropogenic killing, which is also by far the highest hazard overall. We also note that the natural hazard is lower than that for reported poached during the snow season, despite the marginally higher natural incidence, suggesting the latter could be an effect of the interaction of the natural hazard with lowered hazards from other, less prevalent endpoints (e.g., unknown, legal). The higher hazard of poaching (cryptic, through LTF, and reported) relative to other endpoints makes any possible interactions (compensatory or depensatory) among the other hazards (e.g., between natural death and legal killing) seem marginal and possibly influenced by (correlated to) fluctuations in the hazard of poaching. Hence, we caution researchers looking for compensatory or depensatory mechanisms to account for the role of poaching, including its cryptic component, first and foremost.

Our results also indicate different seasonal patterns of hazard for our natural and unknown endpoints, which suggests they should be analyzed separately (contra²⁹). Failure to do so would inflate estimates of anthropogenic mortality and exaggerate the sustainability of lethal management programs that base predictions on estimates of human-caused mortality (e.g.³⁷). Results for endpoints of lower prevalence, such as unknown, collisions, and (to a lesser extent) legal killing when implemented as in Wisconsin (by government agents removing suspected predators of livestock primarily), should be considered preliminary given their respective lower numbers of events per modeled covariate than those recommended to ensure accurate estimation^{38,39}.

The increase in hazard of reported poached and LTF during the hunt/hound/snow season makes this season the deadliest for wolves throughout most of their adult lives (see Supplementary Material Fig. 3). The high hazards of LTF and reported poached, which are higher than all other endpoints for most seasons (hunt/hound, hunt/hound/snow and snow) and throughout t , also confirm poaching as by far the highest mortality hazard for collared adult wolves in Wisconsin throughout their lifetimes^{14,18}.

Furthermore, given attitudes toward wolves became more negative among relevant demographics after wolf hunts were implemented in Wisconsin in 2012³², the general hazard of poaching (cryptic and reported, for all seasons) may have increased relative to our study period (when wolf hunts were not legal) despite possibly resulting in a relatively lower incidence due to the magnitude of the increase in legal killing (e.g., Wisconsin February 2021 wolf hunt⁴⁰). Moreover, the 'facilitated poaching' hypothesis suggests further increases in poaching after permitting wolf hunting, trapping, and hounding (2012–2014, 2021–) relative to only permitting selected legal

Season starts	Season ends	Season (risk_season)
April 15th	June 31st (pre-1991) or July 9th (1991 onward)	'Baseline'; no hounding/hunting/snow (0)
July 1st (pre-1991) OR July 10th (1991 onward)	November 14th	'Hunt/hound' (1)
November 15th	1st Sunday in January	'Hunt/hound/snow' (2)
Monday after 1st Sunday in January	April 14th	'Snow' (3)

Table 2. Intra-wolf-year (April 15th–April 14th) seasons (*risk_season*) characterized by the absence (*baseline* level), presence or overlap of anthropogenic and environmental factors mediating endpoint-specific risk.

killing (our study period)^{17,18,25}. Such an effect of public wolf-hunts would hypothetically be mediated by a policy signal that further devalues wolves or suggests overabundance.

We are not aware of effective efforts by the WDNR to mitigate poaching hazard, neither through increased enforcement nor through public education initiatives. Rather, WDNR efforts have been focused on 'tolerance hunting' through reducing protections, despite multiple lines of evidence pointing to such actions not decreasing and potentially increasing total (cryptic and reported) poaching hazard^{14,18,25,31,32}. In other jurisdictions, such 'tolerance killing' is viewed skeptically as a management tool both scientifically and legally^{13,41–43}. Our results underscore the need for increased protections and anti-poaching interventions to improve the wellbeing of wolves and their populations, and to reduce illegal exploitation of the public trust.

Methods

Data sources and preparation. We analyzed data acquired from the Wisconsin Department of Natural Resources (WDNR) which includes all collared, adult wolves monitored via telemetry (consisting almost entirely of VHF transmitters) in Wisconsin, USA between 1979 and April 14, 2012, published previously in Treves et al.¹⁴ and Santiago-Ávila et al.¹⁸ (n = 495). The dataset includes 487 collared wolves captured and monitored by the WDNR and agents, in addition to 8 wolves initially captured in MI with full monitoring history.

For those wolves monitored until death (n = 242, 49% of monitored individuals), the recorded endpoint classifies their cause of death by one of 5 mutually exclusive causes (following^{14,18}): collision (with vehicles; n = 24, 5%), legal (lethal control by agency personnel; n = 32, 6%), reported poached (illegal killings reported to and evidenced by the agency; n = 88, 18%), natural (unrelated to humans, such as disease or intraspecific strife; n = 77, 16%) or unknown (uncertain cause of death; n = 21, 4%). Dead wolves were recovered via the mortality signal emitted from collars; legal killing by agents; or after reports by private citizens. We defined the date of endpoint for wolves monitored until death as their agency-recorded date of death.

In addition to wolves monitored until death, the data includes 213 wolves (43% of monitored individuals) with a 'lost-to-follow-up' (LTF) endpoint. LTF may occur because: (a) collars stop transmitting (i.e., mechanical failure); (b) permanent migration out of monitoring range; or, (c) cryptic poaching (i.e., concealed and undetected poaching)^{17,18}. The WDNR assigned an LTF endpoint to a wolf if agency personnel was unable to detect the collar signal after various months of aerial or ground telemetry (although effort was not quantified)^{14,18}. We defined the date of endpoint for LTF wolves as the last date of telemetry contact with them. There were 33 LTF wolves (15% of LTF and 7% of collared) that were later recovered, a third of them poached (n = 11). Our main results classify these as LTF, but we include results for a separate endpoint classification of these 33 wolves as 'known-LTF' in supplementary materials. We censored those individuals that survived until the end of the monitoring period (April 15, 2012, n = 40). Our LTF endpoint is conservative given we censored, rather than impute (as in Santiago-Ávila et al. 2020), the fates of n = 26 wolves that disappeared sometime between December 31st, 2011 and April 14th, 2012 and lacked subsequent monitoring or endpoint data in reports between 2012 and 2013 (see Supp Data S2 in Ref.¹⁸). Simulations suggest at least some of these latter wolves may have gone LTF in the winter of 2011–2012¹⁸.

We include two external time-dependent covariates in our statistical models (see below), which are variables that change value at specific dates due to external events, such as a change in season or policy. To include those variables, we split each wolf's monitoring history into time intervals at each specific date of change of that variable so that its value remains constant for each interval. Therefore, each time interval reflects the type of period each wolf was exposed to, and the specific dates during which s/he was exposed.

Our main covariate of interest, *risk_season*, is a four-level categorical variable defining intra-wolf-year periods (wolf-year = April 15th to April 14th) characterized by specific anthropogenic (i.e., hounding and hunting seasons for deer and black bear) and environmental (i.e., snow cover) factors, their overlap, and absence (Table 2). We used specific dates to split each wolf-year in our study period (1979–2012) into four distinct seasons. Our baseline period (*risk_season* = 0) refers to April 15th to June 30th (or to July 9th from 1991 to 2012) and is characterized by the absence of the anthropogenic and environmental conditions present in the other variable levels (i.e., no hounding, no white-tailed deer or black bear hunting, no snow cover). Our hounding and hunting season without snow cover (*risk_season* = 1, 'hunt/hound'), runs from July 1st (July 10th from 1991 to 2012) to Nov 14th. In WI, use of hounds for bear hunting was legalized in 1963 and bear dog training was allowed starting July (1st or 10th) until August 31st. Deer and bear seasons start soon thereafter, in early to mid-September, with the deer season running through the first Sunday in January for most counties (in some counties, the deer season extends to January 31st). Our hounding and hunting season with snow cover (*risk_season* = 2, 'hunt/hound/snow') starts Nov 15th and runs through the first Sunday in January, when deer hunting season ends for most

counties in WI. Average annual duration of snow cover extends to > 140 days along Lake Superior (<http://aos.wisc.edu/>), and most occupied wolf range is in northern Wisconsin. To this data, we added statewide monthly average snowfall (1975–2011) from the WI State Climatology Office, modeling snow cover to include months with an average snow cover of > 1 inch (November through May). Considering both data sources, starting the period on November 15th (average 5.31 in; October, 0.63 in) allowed us to model 151 days of snow cover up to April 14th (average 2.88 in; May, 0.19 in), the end of the wolf-year. Lastly, our snow cover season without hounding or hunting (risk_season = 3, 'snow') runs from the Monday after the first Sunday in January (when deer season closes for most WI counties), until April 14th, as per our snow cover modeling. A breakdown of events per endpoint and time at risk by season is provided in Supplementary Material Table 2.

We also model policy protection periods following Santiago-Ávila et al.¹⁸ (lib_kill, where 1 = reduced protections, i.e., liberalized killing; and 0 = full protection), and include it as a stratifying variable in our statistical models, given evidence of endpoint-specific and sometimes non-proportional effects. In WI, gray wolves were exposed exclusively to full protections under the Endangered Species Act (ESA) from 1979 to March 31, 2003. From April 2003 to 2012, wolves in WI (and MI) were exposed to 11 alternating, sequential and mutually exclusive periods of reduced and restored protections that liberalized and restricted wolf-killing, respectively (Ref.⁴⁴, Supplementary Material Table 2). Periods of reduced protections and liberalized killing (including periods during which permits for 'take' were issued, as well as periods of 'down-' and 'de-listing' from the ESA) were characterized by an announcement of policy change reducing constraints for managers or landowners to kill wolves in response to perceived or actual conflicts, most notably wolf predation on domestic animals.

Statistical tests. Our methods exploit the survival history of collared, monitored adult wolves, measured in days (t), from date of capture and collaring to date of endpoint (i.e., death by multiple causes (see "Data sources and preparation" section) or disappearance). Survival analysis estimates 'time-to-event' functions; i.e., the probability of observing a time interval (T), from beginning of monitoring to endpoint, greater than some stated value ' t '; $S(t) = P(T > t)$. Such techniques also allow for estimating (endpoint-specific) hazard functions, $h_k(t)$; the instantaneous rate of occurrence of an endpoint (k) conditional on not experiencing any endpoint until that time^{45–47}. Semi-parametric, Cox proportional hazard models allow for the estimation of how endpoint-specific hazards change as a function of survival (i.e., monitoring) time and a set of covariates $S(t) = e^{-h_k(t, x, \beta)}$, where x refers to a vector of covariates and β to its parameter estimates. Cox models estimate these covariate effects on endpoint-specific hazard(s) as $h_k(t) = h_{0k}(t)e^{(\beta_1 x_1 + \dots + \beta_j x_j)}$, where $h_{0k}(t)$ is an unestimated baseline hazard function (i.e., semi-parametric), and β_j represent the estimates of HRs for each covariate x_j (HR > 1 is interpreted as an increase, and HR < 1 as a decrease, in hazard).

We employed the Lunn and McNeil⁴⁸ data augmentation technique (by k endpoints) to build stratified (by endpoint) joint Cox proportional hazard models to simultaneously estimate endpoint-specific changes in HRs for each endpoint-season interaction. In using a Cox model, we assume that the endpoint and time-to-endpoint for each wolf is independent of other wolves' (i.e., one wolf's monitoring history and endpoint does not inform others). Because we split the monitoring history of wolves into 'spells' for inclusion of time-dependent covariates (see "Data sources and preparation" section), we cluster analyses by following⁴⁹. We also assume censoring is independent of other endpoints, as we explicitly account for LTF as a separate endpoint given evidence it contains an unaccounted-for source of mortality^{14,17,18,29}. We evaluate compliance with our proportionality assumptions using Schoenfeld residuals^{46,47,50}. We control for non-proportionality of endpoint-season interactions, when necessary, through the inclusion of time-varying coefficients (tvc) for the respective interaction(s). A tvc is an interaction of a parameter with a function of analysis time (t), in our case, $\ln(t)$, to model the change in the main endpoint-season parameter's effect over time. We selected the preferred Cox model considering Akaike's Information Criterion (AIC) and weights, Bayesian Information Criterion (BIC), and compliance with Cox model assumptions.

We then proceed with a competing risk approach by using endpoint-season specific parameter estimates from the best stratified joint Cox model to construct cumulative incidence curves (CIFs) for each endpoint and season. Competing risk approaches focus on the estimation of endpoint-specific CIFs, defined by the failure probability $\text{Prob}(T \leq t, D = k)$; i.e., the cumulative probability of an endpoint, k , occurring over time in the presence of all other competing endpoints^{45,51,52}. These analyses account for the CIF of any endpoint being a function of all other endpoint-specific hazards, $h_k(t)$, thus accounting for the rate of occurrence of that endpoint in addition to how other endpoints influence it⁵³. Thus, joint analysis of hazards and incidence is essential for discerning interactions between endpoint hazards and how they are reflected on each endpoint's incidence.

Consistent with rigorous approaches to competing risk analyses, we present and discuss results for our best performing stratified joint Cox model, by endpoint and season, as well as endpoint-specific CIFs, by season, and synthesize findings^{39,45,51,53}. We conducted all statistical analyses in Stata 16 (StataCorp LLC, College Station, TX, 2019).

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Author contributions

Conceptualization: F.S.A., A.T. Data curation: F.S.A., A.T. Formal analysis: F.S.A. Funding acquisition: F.S.A., A.T. Investigation: F.S.A., A.T. Methodology: F.S.A. Project administration: F.S.A. Resources: F.S.A., A.T. Software: F.S.A. Validation: F.S.A. Writing—original draft: F.S.A. Writing—review and editing: F.S.A., A.T.

Competing interests

FSA declares no competing interests. AT declares no competing interests, and provides his CV (https://faculty.nelson.wisc.edu/treves/archive_BAS/Treves_vita_Jan2020.pdf) and all funding awarded as of 6 Jan 2020 (https://faculty.nelson.wisc.edu/treves/archive_BAS/funding.pdf) for transparency, so readers can decide if they perceive a competing interests.

Additional information

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Poaching of protected wolves fluctuated seasonally and with non-wolf hunting

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Supplementary Tables 1-7:

Endpoint	<i>Baseline</i>	<i>Hunt/hound</i>	<i>Hunt/hound/ snow</i>	<i>Snow</i>	<i>Total</i>
<i>LTF</i>	27	89	34	63	213
<i>legal</i>	9	21	2	0	32
<i>nonhuman</i>	6	20	19	32	77
<i>reported poached</i>	8	20	36	24	88
<i>unknown</i>	4	12	2	3	21
<i>collision</i>	2	14	1	7	24
Total	56	176	94	129	455
Time at risk (t, days)	52,213	116,791	40,225	69,805	279,034

SM Table 1. Number of events (unique wolf IDs) per endpoint and season (*risk_season* variable, see Methods). Wolves that survived to the end of the study period are omitted here and censored at the end of the study period (n=40).

Period start (dd/mm/yyyy)	Period end (dd/mm/yyyy)	Federal status	Policy period** (<i>lib_kill</i>)
15/04/1994	31/03/2003	Listed as endangered	full protections (0)
01/04/2003	30/01/2005	Down-listed to threatened	liberalized killing (1)
31/01/2005	31/03/2005	Relisted	full protections (0)
01/04/2005	13/09/2005	Sub-permit for killing issued	liberalized killing (1)
14/09/2005	23/04/2006	Sub-permit rescinded	full protections (0)
24/04/2006*	31/07/2006	Sub-permit for killing issued	liberalized killing (1)
01/08/2006	11/03/2007	Sub-permit rescinded	full protections (0)
12/03/2007	28/09/2008	Delisted	liberalized killing (1)
29/09/2008	03/05/2009	Relisted	full protections (0)
04/05/2009	30/06/2009	Delisted	liberalized killing (1)
01/07/2009	26/01/2012	Relisted	full protections (0)
27/01/2012	14/04/2012	Delisted	liberalized killing (1)

SM Table 2. Periods of wolf policy changes in Wisconsin and Michigan during our study period (*lib_kill* variable), by policy period in Wisconsin, derived from Refsnider (2009), ESA sec. 4 10(a)(1)(A) and Humane Society of the U.S. et al. v. Jewell (U.S. District Court, D.C., 5 1:13-cv-00186-BAH Document 52, 2014). *WI and MI are identical except for the sub-permit issuance on 6 May 2006 to Michigan instead of issuance on 24 April 2006 to Wisconsin. **Killing a wolf that posed a threat to human safety was always allowed under ESA sec.11(a)(3).

Model selection criteria	Model 1	Model 2	Model 3 (BEST)
<i>Log likelihood</i>	-2114.944	-2106.737	-2107.673
<i>AIC</i>	4263.888	4257.474	4257.347
<i>relative likelihood</i>	6.541	0.127	0.000
<i>AIC weights</i>	0.038	0.938	1.000
<i>BIC</i>	4401.188	4435.157	4426.954
<i>Model description</i>	no tvcs	Model 1 + potential tvcs	Model 2 excluding <i>hunt/hound*lf</i> tvcs

SM Table 3. Model selection statistics for stratified joint Cox models for analyses of hazards using Lunn & McNeil’s data augmentation Method B. We present log-likelihood(1l), AIC and BIC statistics for all models (1-3), along with a brief model description. Model statistics along with significance of tvc parameters (see Table 2 and SM Tables 4-6) suggest Model 3 was the best and most parsimonious model.

SEASON	<i>HUNT/HOUND</i>			<i>HUNT/HOUND/SNOW</i>			<i>SNOW</i>		
	HR (se)	95 CI		HR (se)	95 CI		HR (se)	95 CI	
<i>LTF</i>	1.18 (-0.3)	0.72	1.95	1.19 (-0.37)	0.65	2.19	1.52 (-0.4)	0.90	2.55
<i>legal</i>	1.78 (-0.75)	0.79	4.05	0.72 (-0.69)	0.11	4.71	0.00 (.)	0.00	0.00
<i>reported poached</i>	1.23 (-0.54)	0.52	2.91	7.58*** (-3.34)	3.19	17.99	3.27*** (-1.46)	1.36	7.86
<i>nonhuman</i>	1.34 (-0.68)	0.50	3.62	3.69** (-2.1)	1.21	11.25	3.65** (-2.25)	1.09	12.23
<i>unknown</i>	1.01 (-0.65)	0.29	3.56	0.77 (-0.94)	0.07	8.46	1.01 (-0.71)	0.25	4.03
<i>collision</i>	2.61 (-2.24)	0.49	14.06	0.23 (-0.3)	0.02	3.00	1.17 (-1.09)	0.19	7.27

SM Table 4. Hazard ratio (HR) point estimates from the stratified (by endpoint and protection period) joint Cox Model 1 (our initial model, without tvcs) for n=495 monitored adult wolves, by endpoint and season. We present HRs and compatibility intervals (95 CI) for all endpoint-season interactions relative to a baseline season. Note: * p<0.10, ** p<.05, *** p<0.01.

Parameter (endpoint per season)	rho	chi2	df	Prob>c ²
HUNT/HOUND				
<i>LTF</i>	-0.07194	2.38	1	0.123
<i>legal</i>	0.05599	1.31	1	0.2522
<i>reported poached</i>	0.04518	0.82	1	0.3657
<i>nonhuman</i>	-0.13716	9.94	1	0.0016*
<i>unknown</i>	0.0249	0.44	1	0.5077
<i>collision</i>	-0.03624	0.65	1	0.4198
HUNT/HOUND/SNOW				
<i>LTF</i>	-0.05669	1.61	1	0.2051
<i>legal</i>	-0.04358	0.89	1	0.3443
<i>reported poached</i>	-0.0307	0.47	1	0.491
<i>nonhuman</i>	-0.1185	10.32	1	0.0013*
<i>unknown</i>	-0.07254	6.58	1	0.0103*
<i>collision</i>	0.03395	0.58	1	0.4478
SNOW				
<i>LTF</i>	-0.0423	0.71	1	0.4007
<i>legal</i>	.	.	1	.
<i>reported poached</i>	-0.04843	1.14	1	0.2862
<i>nonhuman</i>	-0.09422	9.62	1	0.0019*
<i>unknown</i>	-0.02747	0.42	1	0.5145
<i>collision</i>	0.0141	0.1	1	0.7577
Global test		34.58	17	0.0071

SM Table 5. Test of proportional hazards assumptions (Chi-squared, χ^2) for parameters in the stratified joint Cox Lunn & McNeil Model 1 (SM Table 4), used for evaluating proportionality assumptions (measured using $\ln(t)$, see Methods). The tests suggest potential non-proportionality present in the tvcs included in Model 2 (identified with a *). The tvcs in Model 3, the preferred stratified joint Cox hazards model, appropriately model the observed non-proportionality through interactions of these endpoint-covariate combinations with analysis time (i.e., tvc) (Table 2).

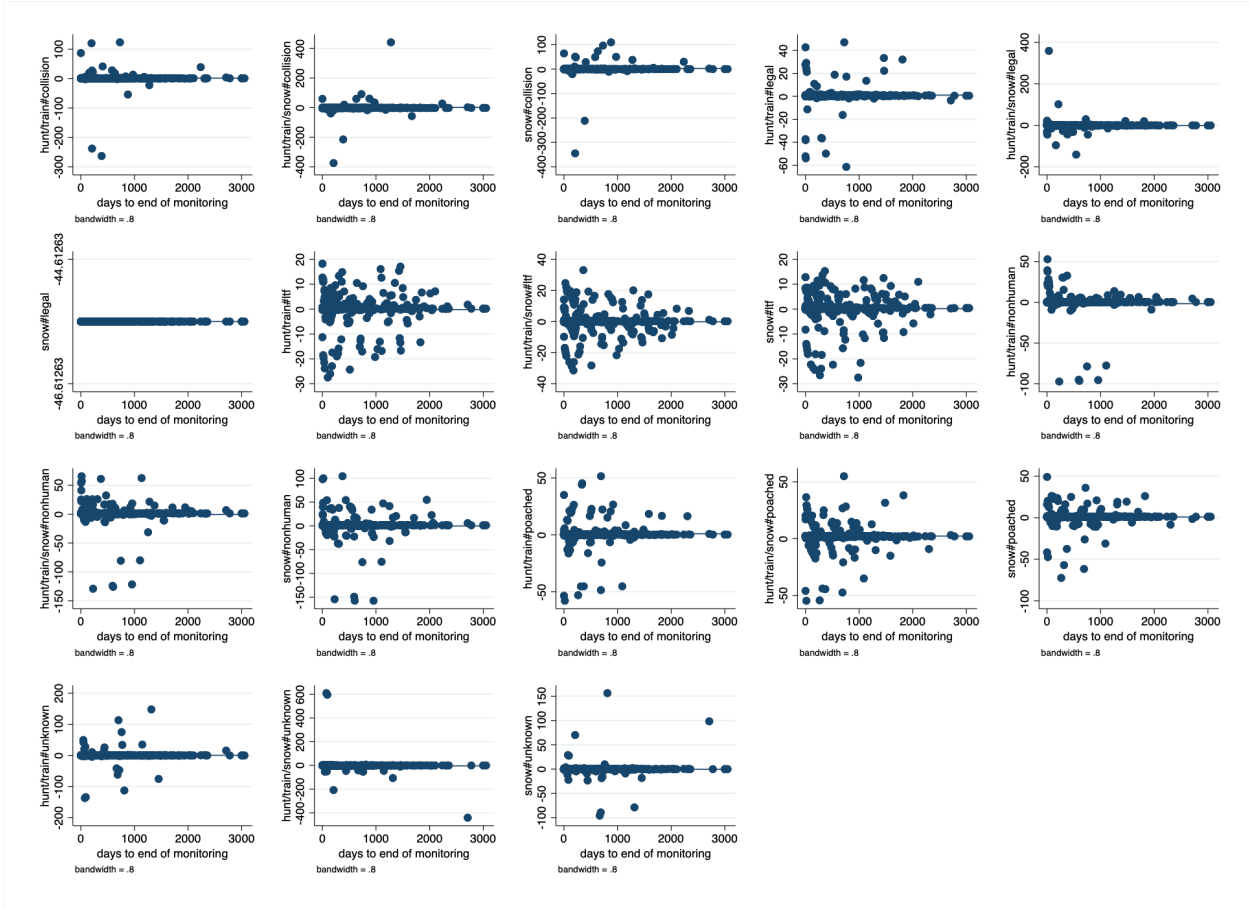
SEASON	<i>HUNT/HOUND</i>			<i>HUNT/HOUND/SNOW</i>			<i>SNOW</i>		
	HR (se)	95 CI		HR (se)	95 CI		HR (se)	95 CI	
<i>LTF</i>	2.48 (-1.44)	0.80	7.71	1.15 (-0.36)	0.62	2.13	1.47 (-0.39)	0.88	2.46
<i>legal</i>	1.78 (-0.75)	0.79	4.05	0.72 (-0.69)	0.11	4.71	0.00 (.)	0.00	0.00
<i>reported poached</i>	1.23 (-0.54)	0.52	2.91	7.58*** (-3.34)	3.19	17.99	3.27*** (-1.46)	1.36	7.86
<i>nonhuman</i>	238.98*** (-405.19)	8.61	6630.66	392.14*** (-709.72)	11.29	13614.75	623.97*** (-1429.7)	7.00	55655.22
<i>unknown</i>	1.20 (-0.8)	0.33	4.44	7671.73*** (-24384.75)	15.11	3894537.72	0.66 (-0.54)	0.13	3.32
<i>collision</i>	2.61 (-2.24)	0.49	14.06	0.23 (-0.3)	0.02	3.00	1.17 (-1.09)	0.19	7.27
<i>tvc - (ln(t))</i>									
<i>LTF</i>	0.86 (-0.09)	0.70	1.06	- -	-	-	- -	-	-
<i>nonhuman</i>	0.39*** (-0.11)	0.23	0.66	0.44*** (-0.13)	0.25	0.77	0.41** (-0.14)	0.20	0.82
<i>unknown</i>	- -	-	-	0.17*** (-0.1)	0.05	0.55	- -	-	-

SM Table 6. Hazard ratio (HR) point estimates from the stratified (by endpoint and protection period) joint Cox Model 2 (including all potential tv_c parameters) for n=495 monitored adult wolves, by endpoint and season. We present HRs and compatibility intervals (95 CI) for all endpoint-season interactions relative to a baseline season. Model 3 drops the LTF tv_c given lack of evidence of non-proportionality. Note: * p<0.10, ** p<.05, *** p<0.01.

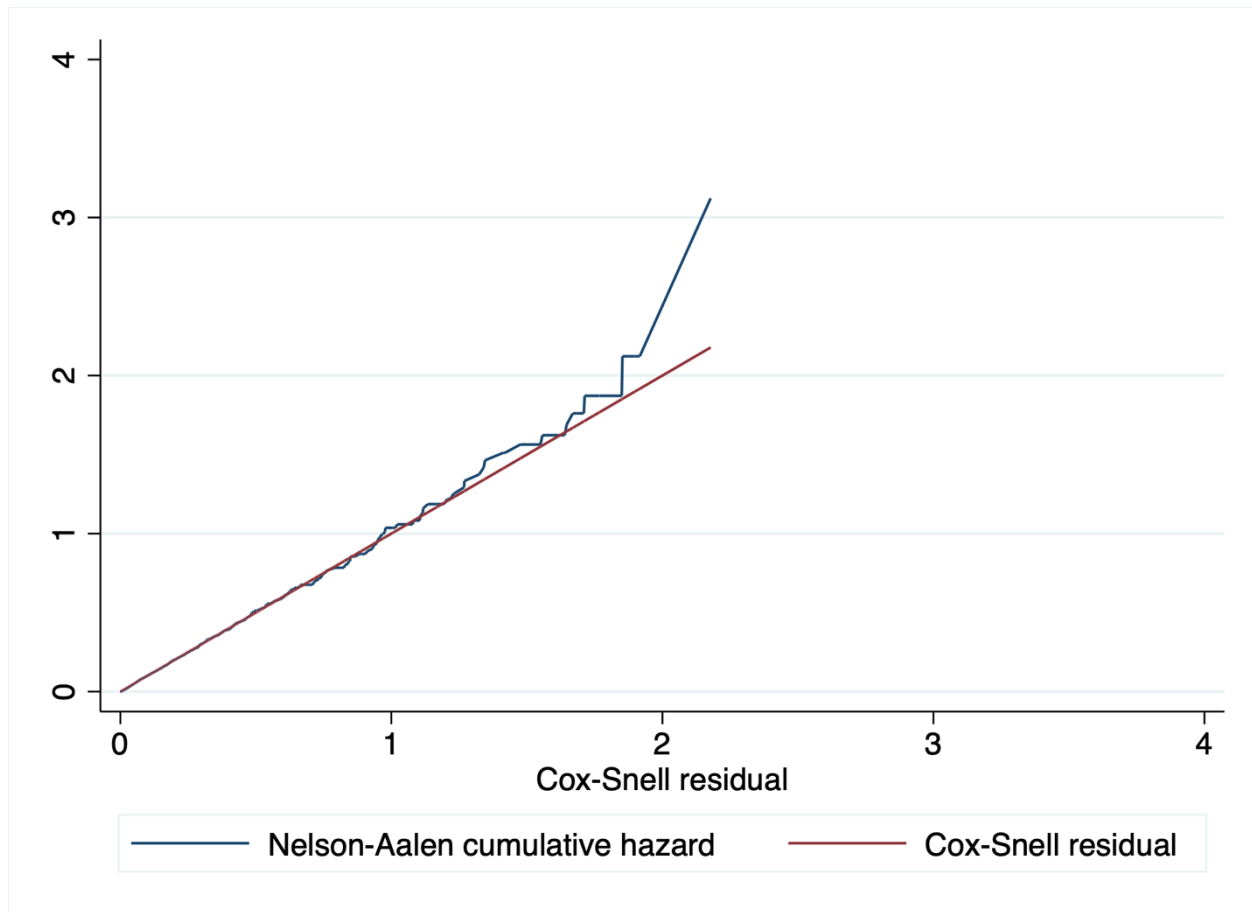
SEASON	HUNT/HOUND			HUNT/HOUND/SNOW			SNOW		
	HR (se)	95 CI		HR (se)	95 CI		HR (se)	95 CI	
<i>LTF</i>	1.25 (-0.35)	0.73	2.16	1.23 (-0.42)	0.63	2.40	1.51 (-0.44)	0.86	2.66
<i>legal</i>	1.78 (-0.75)	0.79	4.05	0.72 (-0.69)	0.11	4.71	0.00 (.)	0.00	0.00
<i>reported poached</i>	1.23 (-0.54)	0.52	2.91	7.58*** (-3.34)	3.19	17.99	3.27*** (-1.46)	1.36	7.86
<i>nonhuman</i>	238.98*** (-405.19)	8.61	6630.66	392.14*** (-709.72)	11.29	13614.75	623.97*** (-1429.7)	7.00	55655.22
<i>unknown</i>	1.20 (-0.8)	0.33	4.44	7671.73*** (-24384.75)	15.11	3894537.72	0.66 (-0.54)	0.13	3.32
<i>collision</i>	2.61 (-2.24)	0.49	14.06	0.23 (-0.3)	0.02	3.00	1.17 (-1.09)	0.19	7.27
<i>known LTF</i>	9.99* (-12.41)	0.87	114.12	0.80 (-0.61)	0.18	3.52	1.25 (-0.82)	0.35	4.50
<i>tvc - (ln(t))</i>									
<i>nonhuman</i>	0.39*** (-0.11)	0.23	0.66	0.44*** (-0.13)	0.25	0.77	0.41** (-0.14)	0.20	0.82
<i>unknown</i>	- -	-	-	0.17*** (-0.1)	0.05	0.55	- -	-	-
<i>known LTF</i>	0.62** (-0.14)	0.40	0.96	- -	-	-	- -	-	-

SM Table 7. Hazard ratio (HR) point estimates from the stratified (by endpoint and protection period) joint Cox Model 3 (Table 2) with n=33 recovered LTF wolves ('known-LTF') reclassified as a separate endpoint (see Methods). We present HRs and compatibility intervals (95 CI) for all endpoint-season interactions relative to a baseline season. Note: * p<0.10, ** p<.05, *** p<0.01.

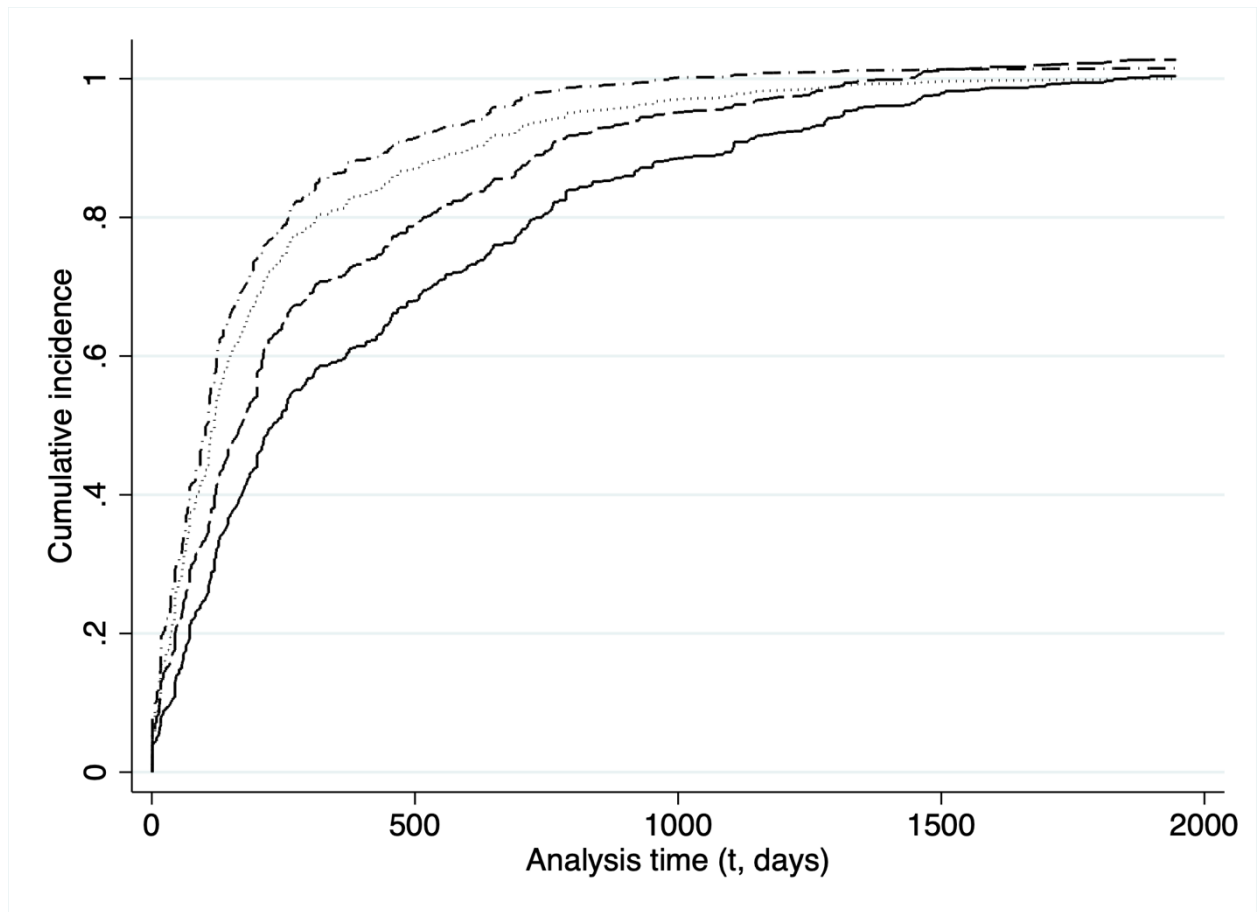
Supplementary Figures 1-3:



SM Figure 1. Schoenfeld residual scatterplots for each endpoint-season combination in Cox Model 1.



SM Figure 2. Cox-Snell generalized residuals (red lines, on x axis) used for evaluating the goodness of fit of stratified joint Cox Model 1. The Nelson-Aalen cumulative hazard (blue lines, on y axis) follows closely the 45° (red) line of Cox-Snell residuals and show overall consistency with the latter, although with some divergence at large values (common when censoring data).



SM Figure 3. Seasonal cumulative incidence curves (CIFs) for all endpoints constructing using all endpoint hazards obtained from our preferred joint stratified Cox model (Model 3, Table 2) for $n=495$ adult monitored wolves in Wisconsin, USA (1979-2012): *baseline* (solid), *hunt/hound* (longdash), *hunt/hound/snow* (dash-dot) and *snow* (dot).

Statistical code for all analyses conducted in STATA

```
*****  
****CREATING VARIABLE FOR MUTUALLY EXCLUSIVE INTRA-YEAR PERIODS****  
use "Supp_Dataset.dta", replace  
  
*stset date_endpoint, failure(cause_endpoint2==2 3 4 5 6 7) exit(failure) origin(time  
capture_date) id(wolf_ID)  
  
rename cause_endpoint2 old_endpoint  
tab old_endpoint cause_endpoint  
encode cause_endpoint, gen(cause_endpoint_enc)  
label var cause_endpoint_enc "coded endpoint"  
order cause_endpoint_enc, after(cause_endpoint)  
tab cause_endpoint cause_endpoint_enc  
  
expand 6  
by wolf_ID, sort: gen cause_endpoint2= _n+1  
order cause_endpoint2, after(cause_endpoint_enc)  
tab cause_endpoint2 cause_endpoint_enc  
  
/*Generating cause of endpoint binary variables*/  
gen collision = cause_endpoint2==2  
gen legal = cause_endpoint2==3  
gen ltf = cause_endpoint2==4  
gen nonhuman = cause_endpoint2==5  
gen poached = cause_endpoint2==6  
gen unknown = cause_endpoint2==7  
gen event = (cause_endpoint2==cause_endpoint_enc)
```

```
gen censored = cause_endpoint_enc==1
```

```
*Generate ID var for each expanded record (wolf_ID - [1-6]) (CLUSTER FOR THIS INSTEAD  
OF wolf_ID)
```

```
gen wolf_ID_exp=wolf_ID+"-"+string(cause_endpoint2, "%02.0f")
```

```
order wolf_ID_exp, after(wolf_ID)
```

```
*Checking stset
```

```
stset date_endpoint, failure(event) exit(failure) origin(time capture_date) id(wolf_ID_exp)
```

```
*****TIME-SPLITTING for time-dependent variables and 'spells'*****
```

```
stsplitt year_split, at(7129 7258 7310 7410 7495 7624 7674 7775 7860 7989 8038 8140 8225  
8354 8402 8505 8590 8719 8773 8871 8956 9085 9137 9236 9321 9450 9501 9601 9686 9815  
9865 9966 10051 10180 10229 10332 10417 10546 10600 10697 10782 10911 10964 11062  
11147 11276 11328 11427 11503 11641 11692 11793 11869 12007 12056 12158 12234 12372  
12420 12523 12599 12737 12791 12888 12964 13102 13155 13254 13330 13468 13519 13619  
13695 13833 13883 13984 14060 14198 14247 14349 14425 14563 14611 14715 14791 14929  
14982 15080 15156 15294 15346 15445 15521 15659 15710 15810 15886 16024 16074 16176  
16252 16390 16438 16541 16617 16755 16809 16906 16982 17120 17173 17271 17347 17485  
17537 17637 17713 17851 17901 18002 18078 18216 18265 18367 18443 18581 18629 18732  
18808 18946 18992 19095 19098) after(capture_date==1/1/1960)
```

```
**Generating intra-year periods time-dep binary variable
```

```
gen risk_season = 0
```

```
label var risk_season "intra-year periods"
```

```
replace risk_season = 1 if year_split==7129 | year_split==7495 | year_split==7860 |  
year_split==8225 | year_split==8590 | year_split==8956 | year_split==9321 |  
year_split==9686 | year_split==10051 | year_split==10417 | year_split==10782 |  
year_split==11147 | year_split==11503 | year_split==11869 | year_split==12234 |  
year_split==12599 | year_split==12964 | year_split==13330 | year_split==13695 |
```

```
year_split==14060 | year_split==14425 | year_split==14791 | year_split==15156 |  
year_split==15521 | year_split==15886 | year_split==16252 | year_split==16617 |  
year_split==16982 | year_split==17347 | year_split==17713 | year_split==18078 |  
year_split==18443 | year_split==18808 | year_split==19098
```

```
replace risk_season = 2 if year_split==7258 | year_split==7624 | year_split==7989 |  
year_split==8354 | year_split==8719 | year_split==9085 | year_split==9450 |  
year_split==9815 | year_split==10180 | year_split==10546 | year_split==10911 |  
year_split==11276 | year_split==11641 | year_split==12007 | year_split==12372 |  
year_split==12737 | year_split==13102 | year_split==13468 | year_split==13833 |  
year_split==14198 | year_split==14563 | year_split==14929 | year_split==15294 |  
year_split==15659 | year_split==16024 | year_split==16390 | year_split==16755 |  
year_split==17120 | year_split==17485 | year_split==17851 | year_split==18216 |  
year_split==18581 | year_split==18946
```

```
replace risk_season = 3 if year_split==6946 | year_split==7310 | year_split==7674 |  
year_split==8038 | year_split==8402 | year_split==8773 | year_split==9137 |  
year_split==9501 | year_split==9865 | year_split==10229 | year_split==10600 |  
year_split==10964 | year_split==11328 | year_split==11692 | year_split==12056 |  
year_split==12420 | year_split==12791 | year_split==13155 | year_split==13519 |  
year_split==13883 | year_split==14247 | year_split==14611 | year_split==14982 |  
year_split==15346 | year_split==15710 | year_split==16074 | year_split==16438 |  
year_split==16809 | year_split==17173 | year_split==17537 | year_split==17901 |  
year_split==18265 | year_split==18629 | year_split==18992
```

```
label define risk_season1 0 "NONE" 1 "hunt/hound" 2 "hunt/hound/snow" 3 "snow"  
label values risk_season risk_season1
```

```
tab risk_season
```

```
tab risk_season if _d==1
```

```
tab cause_endpoint_enc risk_season if _d==1
```

*****PROTECTION PERIOD VARIABLE*****

```
stsplit treat_split, at(15795 16466 16526 16692 16914 17013 17236 17803 18020 18078 19018  
19097) after(capture_date==1/1/1960)
```

***Generating liberalized killing treatment binary variable (1 if lib kill period)

```
gen lib_kill = 0  
replace lib_kill = 1 if treat_split==15795 | treat_split==16526 | treat_split==16914 |  
treat_split==17236 ///  
| treat_split==18020 | treat_split==19018  
order lib_kill treat_split, after(cause_endpoint_agg)
```

```
tab lib_kill if _d==1  
tab cause_endpoint2 lib_kill if _d==1  
tab risk_season lib_kill if _d==1
```

*censoring all other created spells

```
replace event=0 if event==.
```

SAVE DATASET

```
save "Supp_Dataset_expanded.dta", replace
```

```
tab cause_endpoint_enc if _d==1
```

```
*****  
***JOINT ST COX MODELS FOR ALL CAUSE-SPECIFIC HAZARDS*****  
*****
```



```

*Checking stset by wolf_ID_exp (for multiple records)
stset date_endpoint, failure(event) exit(failure) origin(time capture_date) id(wolf_ID_exp)
stdes
stsum

```

```

*****

```

```

***CAUSE-SPECIFIC HAZARD RATES FOR ALL CAUSES SIMULTANEOUSLY***

```

```

*Fit all in same model (basically same results as with separate PH models)

```

```

stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
2.risk_season#2.cause_endpoint2 2.risk_season#3.cause_endpoint2
2.risk_season#4.cause_endpoint2 2.risk_season#5.cause_endpoint2
2.risk_season#6.cause_endpoint2 2.risk_season#7.cause_endpoint2
3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2, efron
strata(cause_endpoint2 lib_kill) robust cluster(wolf_ID)
estimates store fullperiod_allLTFstrat
estat ic

```

```

/*checking assumptions*/

```

```

estat phtest, log detail
estat phtest, plot(1.risk_season#4.cause_endpoint2)
estat phtest, plot(2.risk_season#4.cause_endpoint2)
estat phtest, plot(1.risk_season#5.cause_endpoint2)
estat phtest, plot(2.risk_season#5.cause_endpoint2)
estat phtest, plot(3.risk_season#5.cause_endpoint2)
estat phtest, plot(2.risk_season#7.cause_endpoint2)

```

```

stphplot if cause_endpoint2==2, by(risk_season) nolnt
stphplot if cause_endpoint2==3, by(risk_season) nolnt
stphplot if cause_endpoint2==4, by(risk_season) nolnt
stphplot if cause_endpoint2==5, by(risk_season) nolnt
stphplot if cause_endpoint2==6, by(risk_season) nolnt
stphplot if cause_endpoint2==7, by(risk_season) nolnt

```

*WITH TVCs

```

stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
2.risk_season#2.cause_endpoint2 2.risk_season#3.cause_endpoint2
2.risk_season#4.cause_endpoint2 2.risk_season#5.cause_endpoint2
2.risk_season#6.cause_endpoint2 2.risk_season#7.cause_endpoint2
3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2,
tv(1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
2.risk_season#5.cause_endpoint2 2.risk_season#7.cause_endpoint2
3.risk_season#5.cause_endpoint2) texp(ln(_t)) efron strata(cause_endpoint2 lib_kill) robust
cluster(wolf_ID)
estimates store fullperiod_allLFTVCstrat
estat ic

```

```

stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2 ///
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
2.risk_season#2.cause_endpoint2 2.risk_season#3.cause_endpoint2
2.risk_season#4.cause_endpoint2 2.risk_season#5.cause_endpoint2
2.risk_season#6.cause_endpoint2 2.risk_season#7.cause_endpoint2

```

```

3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2,
tvc(1.risk_season#5.cause_endpoint2 ///
2.risk_season#5.cause_endpoint2 2.risk_season#7.cause_endpoint2
3.risk_season#5.cause_endpoint2) texp(ln(_t)) efron strata(cause_endpoint2 lib_kill) robust
cluster(wolf_ID)
estimates store fullperiod_allLTFTVCstratV2
estat ic

estout fullperiod_allLTFTVCstratV2, eform cells("b(star fmt(3)) ci_1 ci_u" se(par fmt(2)))
starlevels(* 0.10 ** .05 *** 0.01) stats(ll aic bic, labels("Log likelihood" "AIC" "BIC")) legend

```

***DIAGNOSTICS FOR STRATIFIED MODEL**

```

stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
2.risk_season#2.cause_endpoint2 2.risk_season#3.cause_endpoint2
2.risk_season#4.cause_endpoint2 2.risk_season#5.cause_endpoint2
2.risk_season#6.cause_endpoint2 2.risk_season#7.cause_endpoint2
3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2, efron
strata(cause_endpoint2 lib_kill) robust cluster(wolf_ID)

estat phtest, plot(1.risk_season#2.cause_endpoint2) title("") ytitle("hunt/hound#collision")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_COL", replace

```

```
estat phtest, plot(2.risk_season#2.cause_endpoint2) title("") ytitle("hunt/hound/snow#collision")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_COL", replace
estat phtest, plot(3.risk_season#2.cause_endpoint2) title("") ytitle("snow#collision") xtitle("days
to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_COL", replace
```

```
estat phtest, plot(1.risk_season#3.cause_endpoint2) title("") ytitle("hunt/hound#legal")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_LEG", replace
estat phtest, plot(2.risk_season#3.cause_endpoint2) title("") ytitle("hunt/hound/snow#legal")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_LEG", replace
estat phtest, plot(3.risk_season#3.cause_endpoint2) title("") ytitle("snow#legal") xtitle("days to
end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_LEG", replace
```

```
estat phtest, plot(1.risk_season#4.cause_endpoint2) title("") ytitle("hunt/hound#ltf") xtitle("days
to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_LTF", replace
estat phtest, plot(2.risk_season#4.cause_endpoint2) title("") ytitle("hunt/hound/snow#ltf")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_LTF", replace
estat phtest, plot(3.risk_season#4.cause_endpoint2) title("") ytitle("snow#ltf") xtitle("days to end
of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_LTF", replace
```

```
estat phtest, plot(1.risk_season#5.cause_endpoint2) title("") ytitle("hunt/hound#nonhuman")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_NON", replace
```

```

estat phtest, plot(2.risk_season#5.cause_endpoint2) title("")
ytitle("hunt/hound/snow#nonhuman") xtitle("days to end of monitoring")
graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_NON", replace
estat phtest, plot(3.risk_season#5.cause_endpoint2) title("") ytitle("snow#nonhuman")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_NON", replace

estat phtest, plot(1.risk_season#6.cause_endpoint2) title("") ytitle("hunt/hound#poached")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_POA", replace
estat phtest, plot(2.risk_season#6.cause_endpoint2) title("") ytitle("hunt/hound/snow#poached")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_POA", replace
estat phtest, plot(3.risk_season#6.cause_endpoint2) title("") ytitle("snow#poached") xtitle("days
to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_POA", replace

estat phtest, plot(1.risk_season#7.cause_endpoint2) title("") ytitle("hunt/hound#unknown")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period1_UNK", replace
estat phtest, plot(2.risk_season#7.cause_endpoint2) title("") ytitle("hunt/hound/snow#unknown")
xtitle("days to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period2_UNK", replace
estat phtest, plot(3.risk_season#7.cause_endpoint2) title("") ytitle("snow#unknown") xtitle("days
to end of monitoring") graphregion(color(white))
graph save "STCOX_JOINT_Schoenfeld_period3_UNK", replace

graph combine "STCOX_JOINT_Schoenfeld_period1_COL"
"STCOX_JOINT_Schoenfeld_period2_COL" "STCOX_JOINT_Schoenfeld_period3_COL"
"STCOX_JOINT_Schoenfeld_period1_LEG" "STCOX_JOINT_Schoenfeld_period2_LEG"

```

```

"STCOX_JOINT_Schoenfeld_period3_LEG" "STCOX_JOINT_Schoenfeld_period1_LTF"
"STCOX_JOINT_Schoenfeld_period2_LTF" "STCOX_JOINT_Schoenfeld_period3_LTF"
"STCOX_JOINT_Schoenfeld_period1_NON" "STCOX_JOINT_Schoenfeld_period2_NON"
"STCOX_JOINT_Schoenfeld_period3_NON" "STCOX_JOINT_Schoenfeld_period1_POA"
"STCOX_JOINT_Schoenfeld_period2_POA" "STCOX_JOINT_Schoenfeld_period3_POA"
"STCOX_JOINT_Schoenfeld_period1_UNK" "STCOX_JOINT_Schoenfeld_period2_UNK"
"STCOX_JOINT_Schoenfeld_period3_UNK", com graphregion(color(white))
saving("STCOX_JOINT_Schoenfeld_ALL", replace)
graph export "STCOX_JOINT_Schoenfeld_ALL.pdf", replace

```

```

*****stphplots*****

```

```

stphplot if cause_endpoint2==2, by(risk_season) title("Collision") ytitle("-ln{-ln(survival)}")
xtitle("days to end of monitoring") graphregion(color(white)) legend(order(1 "Baseline" 2
"hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none))) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_COL", replace

```

```

stphplot if cause_endpoint2==2, by(risk_season) title("Collision") ytitle("-ln{-ln(survival)}")
xtitle("days to end of monitoring") graphregion(color(white)) legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_COL_nolegend", replace

```

```

stphplot if cause_endpoint2==3, by(risk_season) title("Legal") ytitle("-ln{-ln(survival)}")
xtitle("days to end of monitoring") graphregion(color(white)) legend(order(1 "Baseline" 2
"hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none))) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_LEG", replace

```

```

stphplot if cause_endpoint2==3, by(risk_season) title("Legal") ytitle("-ln{-ln(survival)}")
xtitle("days to end of monitoring") graphregion(color(white)) legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_LEG_nolegend", replace

```

```

stphplot if cause_endpoint2==4, by(risk_season) title("Lost-to-follow-up (LTF)")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))

```

```

legend(order(1 "Baseline" 2 "hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none)))
nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_LTF", replace
stphplot if cause_endpoint2==4, by(risk_season) title("Lost-to-follow-up (LTF)")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_LTF_nolegend", replace

stphplot if cause_endpoint2==5, by(risk_season) title("Nonhuman (natural)")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(order(1 "Baseline" 2 "hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none)))
nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_NON", replace
stphplot if cause_endpoint2==5, by(risk_season) title("Nonhuman (natural)")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_NON_nolegend", replace

stphplot if cause_endpoint2==6, by(risk_season) title("Reported poached")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(order(1 "Baseline" 2 "hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none)))
nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_POA", replace
stphplot if cause_endpoint2==6, by(risk_season) title("Reported poached")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_POA_nolegend", replace

stphplot if cause_endpoint2==7, by(risk_season) title("Unknown cause")
ytitle("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))

```

```

legend(order(1 "Baseline" 2 "hunt/hound" 3 "hunt/hound/snow" 4 "snow") region(lwidth(none)))
nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_UNK", replace
stphplot if cause_endpoint2==7, by(risk_season) title("Unknown cause")
ytile("-ln{-ln(survival)}") xtitle("days to end of monitoring") graphregion(color(white))
legend(off) nolnt
graph save "STCOX_JOINT_InSurvYearPeriod_UNK_nolegend", replace

```

```

graph combine "STCOX_JOINT_InSurvYearPeriod_COL"
"STCOX_JOINT_InSurvYearPeriod_LEG" "STCOX_JOINT_InSurvYearPeriod_LTF"
"STCOX_JOINT_InSurvYearPeriod_NON" "STCOX_JOINT_InSurvYearPeriod_POA"
"STCOX_JOINT_InSurvYearPeriod_UNK", com graphregion(color(white))
saving("STCOX_JOINT_InSurvYearPeriod_ByCauseEnd", replace)
graph export "STCOX_JOINT_InSurvLibKill_ByCauseEnd.pdf", replace

```

Goodness of fit*

```

predict cs, csnell
stset cs, id(wolf_ID_exp) failure(event)
sts generate H = na
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4) graphregion(color(white))
legend(region(lwidth(none)))
graph save "STCOX_JOINT_GoF", replace
graph export "STCOX_JOINT_GoF.pdf", replace
drop cs H

```

```

*****
*****
*****
*****

```

***CHECKING FOR CAUSE-SPECIFIC DIFFERENCES IN HAZARD & PH ASSUMPTION

FOR THE DATASETS USED IN THIS .DO FILE

use "Supp_Dataset.dta", replace

rename cause_endpoint2 old_endpoint

tab old_endpoint cause_endpoint

encode cause_endpoint, gen(cause_endpoint_enc)

label var cause_endpoint_enc "coded endpoint"

order cause_endpoint_enc, after(cause_endpoint)

tab cause_endpoint cause_endpoint_enc

stset date_endpoint, failure(cause_endpoint_enc==2 3 4 5 6 7) exit(failure) origin(time
capture_date) id(wolf_ID)

*****TIME-SPLITTING for time-dependent variables and 'spells'*****

stsplint year_split, at(6946 7044 7120 7258 7310 7410 7486 7624 7674 7775 7851 7989 8038

8140 8216 8354 8402 8505 8581 8719 8773 8871 8947 9085 9137 9236 9312 9450 9501 9601

9677 9815 9865 9966 10042 10180 10229 10332 10408 10546 10600 10697 10773 10911 10964

11062 11138 11276 11328 11427 11503 11641 11692 11793 11869 12007 12056 12158 12234

12372 12420 12523 12599 12737 12791 12888 12964 13102 13155 13254 13330 13468 13519

13619 13695 13833 13883 13984 14060 14198 14247 14349 14425 14563 14611 14715 14791

14929 14982 15080 15156 15294 15346 15445 15521 15659 15710 15810 15886 16024 16074

16176 16252 16390 16438 16541 16617 16755 16809 16906 16982 17120 17173 17271 17347

17485 17537 17637 17713 17851 17901 18002 18078 18216 18265 18367 18443 18581 18629

18732 18808 18946 18992 19095 19098) after(capture_date==1/1/1960)

**Generating snow (1) time-dep binary variable

```

gen risk_season = 0
label var risk_season "intra-year periods"
replace risk_season = 1 if year_split==7120 | year_split==7486 | year_split==7851 |
year_split==8216 | year_split==8581 | year_split==8947 | year_split==9312 | year_split==9677 |
year_split==10042 | year_split==10408 | year_split==10773 | year_split==11138 |
year_split==11503 | year_split==11869 | year_split==12234 | year_split==12599 |
year_split==12964 | year_split==13330 | year_split==13695 | year_split==14060 |
year_split==14425 | year_split==14791 | year_split==15156 | year_split==15521 |
year_split==15886 | year_split==16252 | year_split==16617 | year_split==16982 |
year_split==17347 | year_split==17713 | year_split==18078 | year_split==18443 |
year_split==18808 | year_split==19098

replace risk_season = 2 if year_split==7258 | year_split==7624 | year_split==7989 |
year_split==8354 | year_split==8719 | year_split==9085 | year_split==9450 | year_split==9815 |
year_split==10180 | year_split==10546 | year_split==10911 | year_split==11276 |
year_split==11641 | year_split==12007 | year_split==12372 | year_split==12737 |
year_split==13102 | year_split==13468 | year_split==13833 | year_split==14198 |
year_split==14563 | year_split==14929 | year_split==15294 | year_split==15659 |
year_split==16024 | year_split==16390 | year_split==16755 | year_split==17120 |
year_split==17485 | year_split==17851 | year_split==18216 | year_split==18581 |
year_split==18946

replace risk_season = 3 if year_split==6946 | year_split==7310 | year_split==7674 |
year_split==8038 | year_split==8402 | year_split==8773 | year_split==9137 | year_split==9501 |
year_split==9865 | year_split==10229 | year_split==10600 | year_split==10964 |
year_split==11328 | year_split==11692 | year_split==12056 | year_split==12420 |
year_split==12791 | year_split==13155 | year_split==13519 | year_split==13883 |
year_split==14247 | year_split==14611 | year_split==14982 | year_split==15346 |
year_split==15710 | year_split==16074 | year_split==16438 | year_split==16809 |
year_split==17173 | year_split==17537 | year_split==17901 | year_split==18265 |
year_split==18629 | year_split==18992

```

```

label define risk_seasonl 0 "NONE" 1 "hunt/hound" 2 "hunt/hound/snow" 3 "snow"
label values risk_season risk_seasonl

tab risk_season
tab risk_season if _d==1
tab cause_endpoint_enc risk_season if _d==1

*****PROTECTION PERIOD VARIABLE*****
stsplit treat_split, at(15795 16466 16526 16692 16914 17013 17236 17803 18020 18078 19018
19097) after(capture_date==1/1/1960)

***Generating liberalized killing treatment binary variable (1 if lib kill period)
gen lib_kill = 0
replace lib_kill = 1 if treat_split==15795 | treat_split==16526 | treat_split==16914 |
treat_split==17236 | treat_split==18020 | treat_split==19018
order lib_kill treat_split, after(cause_endpoint_agg)

label define lib_killl 0 "Strict protections" 1 "Reduced protections"
label values lib_kill lib_killl

tab lib_kill if _d==1
tab cause_endpoint_enc lib_kill if _d==1

sort capture_date date_endpoint wolf_ID
gen obs_order = _n
order obs_order, before(wolf_ID)

save "Supp_Dataset_split.dta", replace

```

```
*****  
***ST COX FOR CAUSE-SPECIFIC HAZARDS***  
*****
```

```
use "Supp_Dataset_split.dta", replace
```

```
**COLLISION**
```

```
stset date_endpoint, failure(cause_endpoint_enc==2) exit(failure) origin(time capture_date)  
id(wolf_ID)  
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

```
predict Ch_col_0, basec  
gsort _t -_d  
by _t: replace Ch_col_0 = . if _n > 1  
gen Ch_col_1 = Ch_col_0*exp(_b[1.risk_season])  
gen Ch_col_2 = Ch_col_0*exp(_b[2.risk_season])  
gen Ch_col_3 = Ch_col_0*exp(_b[3.risk_season])
```

```
twoway line Ch_col_* _t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)  
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)  
graphregion(color(white)) legend(off)  
graph save "STCOX_CUMHAZ_COL_STR", replace
```

```
twoway line Ch_col_* _t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)  
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)  
graphregion(color(white)) legend(off)  
graph save "STCOX_CUMHAZ_COL_RED", replace
```

```
graph combine "STCOX_CUMHAZ_COL_STR" "STCOX_CUMHAZ_COL_RED",  
graphregion(color(white)) com saving("STCOX_CUMHAZ_COL_COMB", replace)
```

```
stset date_endpoint, failure(cause_endpoint_enc==2) exit(failure) origin(time capture_date)  
id(wolf_ID)  
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

*predicting HR for CIFs later

```
predict h_col_0, basehc  
gsort _t -_d  
by _t: replace h_col_0 = . if _n > 1  
gen h_col_1 = h_col_0*exp(_b[1.risk_season])  
gen h_col_2 = h_col_0*exp(_b[2.risk_season])  
gen h_col_3 = h_col_0*exp(_b[3.risk_season])
```

```
predict cs, csnell  
stset cs, id(wolf_ID) failure(cause_endpoint_enc==2)  
sts generate H = na  
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)  
drop cs H
```

****LEGAL****

```
stset date_endpoint, failure(cause_endpoint_enc==3) exit(failure) origin(time capture_date)  
id(wolf_ID)  
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

```

predict Ch_leg_0, basec
gsort _t - _d
by _t: replace Ch_leg_0 = . if _n > 1
gen Ch_leg_1 = Ch_leg_0*exp(_b[1.risk_season])
gen Ch_leg_2 = Ch_leg_0*exp(_b[2.risk_season])
gen Ch_leg_3 = Ch_leg_0*exp(_b[3.risk_season])

twoway line Ch_leg_* _t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_LEG_STR", replace

twoway line Ch_leg_* _t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_LEG_RED", replace

graph combine "STCOX_CUMHAZ_LEG_STR" "STCOX_CUMHAZ_LEG_RED",
graphregion(color(white)) com saving("STCOX_CUMHAZ_LEG_COMB", replace)

*predicting HR for CIFs later
predict h_leg_0, basehc
gsort _t - _d
by _t: replace h_leg_0 = . if _n > 1
gen h_leg_1 = h_leg_0*exp(_b[1.risk_season])
gen h_leg_2 = h_leg_0*exp(_b[2.risk_season])
gen h_leg_3 = h_leg_0*exp(_b[3.risk_season])

predict cs, csnell
stset cs, id(wolf_ID) failure(cause_endpoint_enc==3)

```

```
sts generate H = na
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)
drop cs H
```

```
**LTF**
```

```
stset date_endpoint, failure(cause_endpoint_enc==4) exit(failure) origin(time capture_date)
id(wolf_ID)
```

```
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

```
predict Ch_ltf_0, basec
gsort _t - _d
by _t: replace Ch_ltf_0 = . if _n > 1
gen Ch_ltf_1 = Ch_ltf_0*exp(_b[1.risk_season])
gen Ch_ltf_2 = Ch_ltf_0*exp(_b[2.risk_season])
gen Ch_ltf_3 = Ch_ltf_0*exp(_b[3.risk_season])
```

```
twoway line Ch_ltf_* _t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_LTF_STR", replace
```

```
twoway line Ch_ltf_* _t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_LTF_RED", replace
```

```
graph combine "STCOX_CUMHAZ_LTF_STR" "STCOX_CUMHAZ_LTF_RED",  
graphregion(color(white)) com saving("STCOX_CUMHAZ_LTF_COMB", replace)
```

```
*predicting HR for CIFs later
```

```
predict h_ltf_0, basehc  
gsort _t - _d  
by _t: replace h_ltf_0 = . if _n > 1  
gen h_ltf_1 = h_ltf_0*exp(_b[1.risk_season])  
gen h_ltf_2 = h_ltf_0*exp(_b[2.risk_season])  
gen h_ltf_3 = h_ltf_0*exp(_b[3.risk_season])
```

```
predict cs, csnell  
stset cs, id(wolf_ID) failure(cause_endpoint_enc==4)  
sts generate H = na  
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)  
drop cs H
```

```
**POACHED**
```

```
stset date_endpoint, failure(cause_endpoint_enc==6) exit(failure) origin(time capture_date)  
id(wolf_ID)
```

```
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

```
predict Ch_poa_0, basec  
gsort _t - _d
```



```

by _t: replace Ch_poa_0 = . if _n > 1
gen Ch_poa_1 = Ch_poa_0*exp(_b[1.risk_season])
gen Ch_poa_2 = Ch_poa_0*exp(_b[2.risk_season])
gen Ch_poa_3 = Ch_poa_0*exp(_b[3.risk_season])

twoway line Ch_poa_*_t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_POA_STR", replace

twoway line Ch_poa_*_t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_POA_RED", replace

graph combine "STCOX_CUMHAZ_POA_STR" "STCOX_CUMHAZ_POA_RED",
graphregion(color(white)) com saving("STCOX_CUMHAZ_POA_COMB", replace)

*predicting HR for CIFs later
predict h_poa_0, basehc
gsort _t - _d
by _t: replace h_poa_0 = . if _n > 1
gen h_poa_1 = h_poa_0*exp(_b[1.risk_season])
gen h_poa_2 = h_poa_0*exp(_b[2.risk_season])
gen h_poa_3 = h_poa_0*exp(_b[3.risk_season])

predict cs, csnell
stset cs, id(wolf_ID) failure(cause_endpoint_enc==6)
sts generate H = na
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)

```

```
drop cs H
```

```
**UNKNOWN**
```

```
stset date_endpoint, failure(cause_endpoint_enc==7) exit(failure) origin(time capture_date)  
id(wolf_ID)
```

```
stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
```

```
predict Ch_unk_0, basec
```

```
gsort _t -_d
```

```
by _t: replace Ch_unk_0 = . if _n > 1
```

```
gen Ch_unk_1 = Ch_unk_0*exp(_b[1.risk_season])
```

```
gen Ch_unk_2 = Ch_unk_0*exp(_b[2.risk_season])
```

```
gen Ch_unk_3 = Ch_unk_0*exp(_b[3.risk_season])
```

```
twoway line Ch_unk*_t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
```

```
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
```

```
graphregion(color(white)) legend(off)
```

```
graph save "STCOX_CUMHAZ_UNK_STR", replace
```

```
twoway line Ch_unk*_t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
```

```
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
```

```
graphregion(color(white)) legend(off)
```

```
graph save "STCOX_CUMHAZ_UNK_RED", replace
```

```
graph combine "STCOX_CUMHAZ_UNK_STR" "STCOX_CUMHAZ_UNK_RED",
```

```
graphregion(color(white)) com saving("STCOX_CUMHAZ_UNK_COMB", replace)
```

```

*predicting HR for CIFs later
predict h_unk_0, basehc
gsort _t - _d
by _t: replace h_unk_0 = . if _n > 1
gen h_unk_1 = h_unk_0*exp(_b[1.risk_season])
gen h_unk_2 = h_unk_0*exp(_b[2.risk_season])
gen h_unk_3 = h_unk_0*exp(_b[3.risk_season])

predict cs, csnell
stset cs, id(wolf_ID) failure(cause_endpoint_enc==7)
sts generate H = na
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)
drop cs H

**NONHUMAN**
stset date_endpoint, failure(cause_endpoint_enc==5) exit(failure) origin(time capture_date)
id(wolf_ID)

stcox i.risk_season, strata(lib_kill) efron robust cluster(wolf_ID)
stcox i.risk_season, tvc(i.risk_season) texp(ln(_t)) strata(lib_kill) efron robust cluster(wolf_ID)

*for tvcs without the command (to use for predict command below)
stsplit, at(failures)
gen lnt = ln(_t)

stcox i.risk_season 1.risk_season#c.lnt 2.risk_season#c.lnt 3.risk_season#c.lnt, strata(lib_kill)
efron robust cluster(wolf_ID)

```

```
stcox i.risk_season 1.risk_season#c.lnt 2.risk_season#c.lnt 3.risk_season#c.lnt, strata(lib_kill)
efron robust cluster(wolf_ID) nohr
```

```
predict Ch_non_0, basec
gsort _t - _d
by _t: replace Ch_non_0 = . if _n > 1
gen Ch_non_1 = Ch_non_0*exp(_b[1.risk_season]+(-.9459155*lnt))
gen Ch_non_2 = Ch_non_0*exp(_b[2.risk_season]+(-.8267926*lnt))
gen Ch_non_3 = Ch_non_0*exp(_b[3.risk_season]+(-.8997016*lnt))
```

```
twoway line Ch_non_* _t if _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_NON_STR", replace
```

```
twoway line Ch_non_* _t if _d==1 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CUMHAZ_NON_RED", replace
```

```
graph combine "STCOX_CUMHAZ_NON_STR" "STCOX_CUMHAZ_NON_RED", com
saving("STCOX_CUMHAZ_NON_COMB", replace)
```

*predicting HR for CIFs later

```
predict h_non_0, basehc
gsort _t - _d
by _t: replace h_non_0 = . if _n > 1
gen h_non_1 = h_non_0*exp(_b[1.risk_season]+(-.9459155*lnt))
gen h_non_2 = h_non_0*exp(_b[2.risk_season]+(-.8267926*lnt))
gen h_non_3 = h_non_0*exp(_b[3.risk_season]+(-.8997016*lnt))
```

```

predict cs, csnell
stset cs, id(wolf_ID) failure(cause_endpoint_enc==5)
sts generate H = na
line H cs cs, sort xlab(0 1 to 4) ylab(0 1 to 4)
drop cs H

```

```

stset date_endpoint, failure(cause_endpoint_enc==2 3 4 5 6 7) exit(failure) origin(time
capture_date) id(wolf_ID)

```

```

tway line Ch_ltf_0 Ch_ltf_1 Ch_poa_0 Ch_poa_1 Ch_leg_0 Ch_leg_1 Ch_non_0 Ch_non_1
_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair stepstair
stepstair stepstair stepstair)sort lpattern(solid solid longdash longdash shortdash_dot
shortdash_dot dot dot) lcolor(black gray black gray black gray black gray) ytitle("Cumulative
hazard") xtitle("Analysis time (t, days)") graphregion(color(white)) legend(off)
graph save "STCOX_Ch_LTF_POA_LEG_NON_COMB_1_BG_t", replace

```

```

tway line Ch_ltf_0 Ch_ltf_2 Ch_poa_0 Ch_poa_2 Ch_leg_0 Ch_leg_2 Ch_non_0 Ch_non_2
_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair stepstair
stepstair stepstair stepstair)sort lpattern(solid solid longdash longdash shortdash_dot
shortdash_dot dot dot) lcolor(black gray black gray black gray black gray) ytitle("Cumulative
hazard") xtitle("Analysis time (t, days)") graphregion(color(white)) legend(off)
graph save "STCOX_Ch_LTF_POA_LEG_NON_COMB_2_BG_t", replace

```

```

tway line Ch_ltf_0 Ch_ltf_3 Ch_poa_0 Ch_poa_3 Ch_leg_0 Ch_leg_3 Ch_non_0 Ch_non_3
_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair stepstair stepstair stepstair stepstair
stepstair stepstair stepstair)sort lpattern(solid solid longdash longdash shortdash_dot
shortdash_dot dot dot) lcolor(black gray black gray black gray black gray) ytitle("Cumulative
hazard") xtitle("Analysis time (t, days)") graphregion(color(white)) legend(off)

```

```
graph save "STCOX_Ch_LTF_POA_LEG_NON_COMB_3_BG_t", replace
```

```
twoway line Ch_non_*_t
```

```
gen Ch_hts_poa_ltf = Ch_poa_2/Ch_ltf_2
```

```
sum Ch_hts_poa_ltf
```

```
twoway line Ch_hts_poa_ltf_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair) sort  
lpattern(solid) lcolor(black) ytitle("Cumulative hazard") xtitle("Analysis time (t, days)")  
graphregion(color(white)) legend(off)
```

```
drop Ch_hts_poa_ltf
```

```
gen Ch_hts_non_ltf = Ch_non_2/Ch_ltf_2
```

```
sum Ch_hts_non_ltf if _d==1 & lib_kill==0
```

```
twoway line Ch_hts_non_ltf_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair) sort  
lpattern(solid) lcolor(black) ytitle("Cumulative hazard") xtitle("Analysis time (t, days)")  
graphregion(color(white)) legend(off)
```

```
gen Ch_s_non_ltf = Ch_non_3/Ch_ltf_3
```

```
sum Ch_s_non_ltf if _d==1 & lib_kill==0 & _t>=100
```

```
twoway line Ch_s_non_ltf_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair) sort  
lpattern(solid) lcolor(black) ytitle("Cumulative hazard") xtitle("Analysis time (t, days)")  
graphregion(color(white)) legend(off)
```

```
gen Ch_s_poa_ltf = Ch_poa_3/Ch_ltf_3
sum Ch_s_poa_ltf if _d==1 & lib_kill==0 & _t>=100
```

```
twoway line Ch_s_poa_ltf_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair) sort
lpattern(solid) lcolor(black) ytitle("Cumulative hazard") xtitle("Analysis time (t, days)")
graphregion(color(white)) legend(off)
```

```
gen Ch_s_non_poa = Ch_non_3/Ch_poa_3
sum Ch_s_non_poa if _d==1 & lib_kill==0 & _t>=100
```

```
twoway line Ch_s_non_poa_t if _t<=2000 & _d==1 & lib_kill==0, connect(stepstair) sort
lpattern(solid) lcolor(black) ytitle("Cumulative hazard") xtitle("Analysis time (t, days)")
graphregion(color(white)) legend(off)
```

```
*****
```

```
**CREATING CIFs WITH ABOVE HAZARD CONTRIBUTIONS
```

```
stset date_endpoint, failure(cause_endpoint_enc==2 3 4 5 6 7) exit(failure) origin(time
capture_date) id(wolf_ID)
```

```
drop if missing(h_col_0) & missing(h_leg_0) & missing(h_ltf_0) & missing(h_non_0) &
missing(h_poa_0) & missing(h_unk_0)
```

```
replace h_col_0=0 if missing(h_col_0)
```

```
replace h_col_1=0 if missing(h_col_1)
```

```
replace h_col_2=0 if missing(h_col_2)
```

```
replace h_col_3=0 if missing(h_col_3)
```

```
replace h_leg_0=0 if missing(h_leg_0)
```

```
replace h_leg_1=0 if missing(h_leg_1)
```

```
replace h_leg_2=0 if missing(h_leg_2)
```

replace h_leg_3=0 if missing(h_leg_3)

replace h_ltf_0=0 if missing(h_ltf_0)

replace h_ltf_1=0 if missing(h_ltf_1)

replace h_ltf_2=0 if missing(h_ltf_2)

replace h_ltf_3=0 if missing(h_ltf_3)

replace h_non_0=0 if missing(h_non_0)

replace h_non_1=0 if missing(h_non_1)

replace h_non_2=0 if missing(h_non_2)

replace h_non_3=0 if missing(h_non_3)

replace h_poa_0=0 if missing(h_poa_0)

replace h_poa_1=0 if missing(h_poa_1)

replace h_poa_2=0 if missing(h_poa_2)

replace h_poa_3=0 if missing(h_poa_3)

replace h_unk_0=0 if missing(h_unk_0)

replace h_unk_1=0 if missing(h_unk_1)

replace h_unk_2=0 if missing(h_unk_2)

replace h_unk_3=0 if missing(h_unk_3)

****calculating event-free survivor functions**

sort _t

gen S_0 = exp(sum(log(1- h_col_0 - h_leg_0 - h_ltf_0 - h_non_0 - h_poa_0 - h_unk_0))) if
_d==1

gen S_1 = exp(sum(log(1- h_col_1 - h_leg_1 - h_ltf_1 - h_non_1 - h_poa_1 - h_unk_1))) if
_d==1

gen S_2 = exp(sum(log(1- h_col_2 - h_leg_2 - h_ltf_2 - h_non_2 - h_poa_2 - h_unk_2))) if
_d==1


```
gen S_3 = exp(sum(log(1- h_col_3 - h_leg_3 - h_ltf_3 - h_non_3 - h_poa_3 - h_unk_3))) if
_d==1
```

```
twoway line S_*_t if t<=2000, connect(stepstair stepstair stepstair stepstair) sort lpattern(solid
dash_dot dash longdash) lcolor(navy orange maroon green) graphregion(color(white))
legend(off)
graph save "STCOX_CIF_ALL", replace
```

```
twoway line S_*_t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_ALL_STR", replace
```

```
twoway line S_*_t if _t<=500 & lib_kill==1, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_ALL_RED", replace
```

```
graph combine "STCOX_CIF_ALL_STR" "STCOX_CIF_ALL_RED",
graphregion(color(white)) com saving("STCOX_CIF_ALL_COMB", replace)
```

*calculating CIFs

*colissions

```
gen cif_col_0 = sum(S_0[_n-1]*h_col_0)
```

```
gen cif_col_1 = sum(S_1[_n-1]*h_col_1)
```

```
gen cif_col_2 = sum(S_2[_n-1]*h_col_2)
```

```
gen cif_col_3 = sum(S_3[_n-1]*h_col_3)
```

```

tway line cif_col_*_t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_COL", replace

```

```

tway line cif_col_*_t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_COL_STR", replace

```

```

tway line cif_col_*_t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_COL_RED", replace

```

```

graph combine "STCOX_CIF_COL_STR" "STCOX_CIF_COL_RED",
graphregion(color(white)) com saving("STCOX_CIF_COL_COMB", replace)

```

*legal

```

gen cif_leg_0 = sum(S_0[_n-1]*h_leg_0)
gen cif_leg_1 = sum(S_1[_n-1]*h_leg_1)
gen cif_leg_2 = sum(S_2[_n-1]*h_leg_2)
gen cif_leg_3 = sum(S_3[_n-1]*h_leg_3)
tway line cif_leg_*_t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LEG", replace

```

```

tway line cif_leg_* _t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LEG_STR", replace

```

```

tway line cif_leg_* _t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LEG_RED", replace

```

```

graph combine "STCOX_CIF_LEG_STR" "STCOX_CIF_LEG_RED",
graphregion(color(white)) com saving("STCOX_CIF_LEG_COMB", replace)

```

*ltf

```

gen cif_ltf_0 = sum(S_0[_n-1]*h_ltf_0)
gen cif_ltf_1 = sum(S_1[_n-1]*h_ltf_1)
gen cif_ltf_2 = sum(S_2[_n-1]*h_ltf_2)
gen cif_ltf_3 = sum(S_3[_n-1]*h_ltf_3)
tway line cif_ltf_* _t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF", replace

```

```

tway line cif_ltf_* _t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_STR", replace

```

```
twoway line cif_ltf_*_t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_RED", replace
```

```
graph combine "STCOX_CIF_LTF_STR" "STCOX_CIF_LTF_RED",
graphregion(color(white)) com saving("STCOX_CIF_LTF_COMB", replace)
```

*nonhuman

```
gen cif_non_0 = sum(S_0[_n-1]*h_non_0)
gen cif_non_1 = sum(S_1[_n-1]*h_non_1)
gen cif_non_2 = sum(S_2[_n-1]*h_non_2)
gen cif_non_3 = sum(S_3[_n-1]*h_non_3)
twoway line cif_non_*_t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_NON", replace
```

```
twoway line cif_non_*_t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_NON_STR", replace
```

```
twoway line cif_non_*_t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_NON_RED", replace
```

```
graph combine "STCOX_CIF_NON_STR" "STCOX_CIF_NON_RED", com
saving("STCOX_CIF_NON_COMB", replace)
```

```
*poached
```

```
gen cif_poa_0 = sum(S_0[_n-1]*h_poa_0)
gen cif_poa_1 = sum(S_1[_n-1]*h_poa_1)
gen cif_poa_2 = sum(S_2[_n-1]*h_poa_2)
gen cif_poa_3 = sum(S_3[_n-1]*h_poa_3)
twoway line cif_poa_* _t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_POA", replace
```

```
twoway line cif_poa_* _t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_POA_STR", replace
```

```
twoway line cif_poa_* _t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_POA_RED", replace
```

```
graph combine "STCOX_CIF_POA_STR" "STCOX_CIF_POA_RED",
graphregion(color(white)) com saving("STCOX_CIF_POA_COMB", replace)
```

```
*unknown
```

```
gen cif_unk_0 = sum(S_0[_n-1]*h_unk_0)
```

```

gen cif_unk_1 = sum(S_1[_n-1]*h_unk_1)
gen cif_unk_2 = sum(S_2[_n-1]*h_unk_2)
gen cif_unk_3 = sum(S_3[_n-1]*h_unk_3)
twoway line cif_unk_* _t if _t<=300, connect(stepstair stepstair stepstair stepstair) sort
lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_UNK", replace

```

```

twoway line cif_unk_* _t if _t<=300 & lib_kill==0, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_UNK_STR", replace

```

```

twoway line cif_unk_* _t if _t<=300 & lib_kill==1, connect(stepstair stepstair stepstair stepstair)
sort lpattern(solid dash_dot dash longdash) lcolor(navy orange maroon green)
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_UNK_RED", replace

```

```

graph combine "STCOX_CIF_UNK_STR" "STCOX_CIF_UNK_RED",
graphregion(color(white)) com saving("STCOX_CIF_UNK_COMB", replace)

```

```

twoway line cif_ltf_* cif_poa_* cif_leg_* _t, connect(stepstair stepstair stepstair stepstair
stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair) sort lpattern(solid solid
solid solid longdash longdash longdash longdash shortdash_dot shortdash_dot shortdash_dot
shortdash_dot) lcolor(navy orange maroon green navy orange maroon green navy orange maroon
green) graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_COMB", replace

```

```

twoway line cif_ltf_* cif_poa_* cif_non_* _t if _t<=200, connect(stepstair stepstair stepstair
stepstair stepstair stepstair stepstair stepstair stepstair stepstair) sort
lpattern(solid solid solid solid longdash longdash longdash longdash shortdash_dot
shortdash_dot shortdash_dot shortdash_dot) lcolor(navy orange maroon green navy orange
maroon green navy orange maroon green) graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_NON_COMB", replace

```

```

twoway line cif_ltf_0 cif_ltf_1 cif_poa_0 cif_poa_1 cif_leg_0 cif_leg_1 _t if _t<=200,
connect(stepstair stepstair stepstair stepstair stepstair stepstair)sort lpattern(solid solid longdash
longdash shortdash_dot shortdash_dot) lcolor(navy orange navy orange avy orange )
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_COMB", replace

```

```

twoway line cif_ltf_0 cif_ltf_1 cif_poa_0 cif_poa_1 cif_leg_0 cif_leg_1 cif_non_0 cif_non_1 _t
if _t<=200, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(navy orange
navy orange navy orange navy orange) graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_1", replace

```

```

twoway line cif_ltf_0 cif_ltf_2 cif_poa_0 cif_poa_2 cif_leg_0 cif_leg_2 cif_non_0 cif_non_2 _t
if _t<=200, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(navy maroon
navy maroon navy maroon navy maroon) graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_2", replace

```

```

twoway line cif_ltf_0 cif_ltf_3 cif_poa_0 cif_poa_3 cif_leg_0 cif_leg_3 cif_non_0 cif_non_3 _t
if _t<=200, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(navy green
navy green navy green navy green) graphregion(color(white)) legend(off)

```

```
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_3", replace
```

****BLACK/GRAY FIGURES****

```
twoway line cif_ltf_0 cif_ltf_1 cif_poa_0 cif_poa_1 cif_leg_0 cif_leg_1 cif_non_0 cif_non_1 _t
if _t<=2000, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(black gray
black gray black gray black gray) ytitle("Cumulative incidence") xtitle("Analysis time (t, days)")
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_1_BG", replace
```

```
twoway line cif_ltf_0 cif_ltf_2 cif_poa_0 cif_poa_2 cif_leg_0 cif_leg_2 cif_non_0 cif_non_2 _t
if _t<=2000, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(black gray
black gray black gray black gray) ytitle("Cumulative incidence") xtitle("Analysis time (t, days)")
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_2_BG", replace
```

```
twoway line cif_ltf_0 cif_ltf_3 cif_poa_0 cif_poa_3 cif_leg_0 cif_leg_3 cif_non_0 cif_non_3 _t
if _t<=2000, connect(stepstair stepstair stepstair stepstair stepstair stepstair stepstair stepstair)sort
lpattern(solid solid longdash longdash shortdash_dot shortdash_dot dot dot) lcolor(black gray
black gray black gray black gray) ytitle("Cumulative incidence") xtitle("Analysis time (t, days)")
graphregion(color(white)) legend(off)
graph save "STCOX_CIF_LTF_POA_LEG_NON_COMB_3_BG", replace
```

*ALL CIFs

```
gen cif_all_0 = cif_col_0 + cif_leg_0 + cif_ltf_0 + cif_non_0 + cif_poa_0 + cif_unk_0
gen cif_all_1 = cif_col_1 + cif_leg_1 + cif_ltf_1 + cif_non_1 + cif_poa_1 + cif_unk_1
```



```
gen cif_all_2 = cif_col_2 + cif_leg_2 + cif_ltf_2 + cif_non_2 + cif_poa_2 + cif_unk_2
gen cif_all_3 = cif_col_3 + cif_leg_3 + cif_ltf_3 + cif_non_3 + cif_poa_3 + cif_unk_3
```

```
twoway line cif_all_* _t if _t<=2000, connect(stepstair stepstair stepstair stepstair) sort
lcolor(navy orange maroon green) ytitle("Cumulative incidence") xtitle("Analysis time (t,
days)") graphregion(color(white)) legend(off)graphregion(color(white)) legend(off)
graph save "STCOX_CIF_ALL", replace
```

```
*****
*****
*****
*****
*****
```

```
*****
****USING X-LTF DATASET (RECOVERED LTF) FOR CHECKS*****
```

```
**Reclassifying LTF known fates
use "Supp_Dataset.dta", replace
tab collar_status2 cause_death2
list wolf_ID date_endpoint last_contact collar_recovery_death cause_death cause_endpoint if
collar_status2==2 & cause_death2!=.
replace cause_endpoint = "x-ltf" if collar_status2==2 & cause_death2!=.
tab cause_endpoint
save "Supp_Dataset_xLTF.dta", replace
```

```
list wolf_ID date_endpoint cause_endpoint if year_endpoint<=1984 & month_endpoint>=4 &
month_endpoint<=9
```

```
rename cause_endpoint2 old_endpoint
tab old_endpoint cause_endpoint
encode cause_endpoint, gen(cause_endpoint_enc)
label var cause_endpoint_enc "coded endpoint"
order cause_endpoint_enc, after(cause_endpoint)
tab cause_endpoint cause_endpoint_enc
```

```
expand 7
by wolf_ID, sort: gen cause_endpoint2= _n+1
order cause_endpoint2, after(cause_endpoint_enc)
tab cause_endpoint2 cause_endpoint_enc
```

```
/*Generating cause of endpoint binary variables*/
gen collision = cause_endpoint2==2
gen legal = cause_endpoint2==3
gen ltf = cause_endpoint2==4
gen nonhuman = cause_endpoint2==5
gen poached = cause_endpoint2==6
gen unknown = cause_endpoint2==7
gen x_ltf = cause_endpoint2==8
gen event = (cause_endpoint2==cause_endpoint_enc)
gen censored = cause_endpoint_enc==1
```

```
*Generate ID var for each expanded record (wolf_ID - [1-6]) (CLUSTER FOR THIS INSTEAD
OF wolf_ID)
```

```
gen wolf_ID_exp=wolf_ID+"-"+string(cause_endpoint2, "%02.0f")
order wolf_ID_exp, after(wolf_ID)
```

```
*Checking stset
```

```
stset date_endpoint, failure(event) exit(failure) origin(time capture_date) id(wolf_ID_exp)
```

*****TIME-SPLITTING for time-dependent variables and 'spells'*****

```
stsplit year_split, at(7129 7258 7310 7410 7495 7624 7674 7775 7860 7989 8038 8140 8225
8354 8402 8505 8590 8719 8773 8871 8956 9085 9137 9236 9321 9450 9501 9601 9686 9815
9865 9966 10051 10180 10229 10332 10417 10546 10600 10697 10782 10911 10964 11062
11147 11276 11328 11427 11503 11641 11692 11793 11869 12007 12056 12158 12234 12372
12420 12523 12599 12737 12791 12888 12964 13102 13155 13254 13330 13468 13519 13619
13695 13833 13883 13984 14060 14198 14247 14349 14425 14563 14611 14715 14791 14929
14982 15080 15156 15294 15346 15445 15521 15659 15710 15810 15886 16024 16074 16176
16252 16390 16438 16541 16617 16755 16809 16906 16982 17120 17173 17271 17347 17485
17537 17637 17713 17851 17901 18002 18078 18216 18265 18367 18443 18581 18629 18732
18808 18946 18992 19095 19098) after(capture_date==1/1/1960)
```

**Generating intra-year periods time-dep binary variable

```
gen risk_season = 0
```

```
label var risk_season "intra-year periods"
```

```
replace risk_season = 1 if year_split==7129 | year_split==7495 | year_split==7860 |
year_split==8225 | year_split==8590 | year_split==8956 | year_split==9321 |
year_split==9686 | year_split==10051 | year_split==10417 | year_split==10782 |
year_split==11147 | year_split==11503 | year_split==11869 | year_split==12234 |
year_split==12599 | year_split==12964 | year_split==13330 | year_split==13695 |
year_split==14060 | year_split==14425 | year_split==14791 | year_split==15156 |
year_split==15521 | year_split==15886 | year_split==16252 | year_split==16617 |
year_split==16982 | year_split==17347 | year_split==17713 | year_split==18078 |
year_split==18443 | year_split==18808 | year_split==19098
```

```
replace risk_season = 2 if year_split==7258 | year_split==7624 | year_split==7989 |
year_split==8354 | year_split==8719 | year_split==9085 | year_split==9450 |
year_split==9815 | year_split==10180 | year_split==10546 | year_split==10911 |
year_split==11276 | year_split==11641 | year_split==12007 | year_split==12372 |
year_split==12737 | year_split==13102 | year_split==13468 | year_split==13833 |
```

```
year_split==14198 | year_split==14563 | year_split==14929 | year_split==15294 |  
year_split==15659 | year_split==16024 | year_split==16390 | year_split==16755 |  
year_split==17120 | year_split==17485 | year_split==17851 | year_split==18216 |  
year_split==18581 | year_split==18946
```

```
replace risk_season = 3 if year_split==6946 | year_split==7310 | year_split==7674 |  
year_split==8038 | year_split==8402 | year_split==8773 | year_split==9137 |  
year_split==9501 | year_split==9865 | year_split==10229 | year_split==10600 |  
year_split==10964 | year_split==11328 | year_split==11692 | year_split==12056 |  
year_split==12420 | year_split==12791 | year_split==13155 | year_split==13519 |  
year_split==13883 | year_split==14247 | year_split==14611 | year_split==14982 |  
year_split==15346 | year_split==15710 | year_split==16074 | year_split==16438 |  
year_split==16809 | year_split==17173 | year_split==17537 | year_split==17901 |  
year_split==18265 | year_split==18629 | year_split==18992
```

```
label define risk_season1 0 "NONE" 1 "hunt/hound" 2 "hunt/hound/snow" 3 "snow"  
label values risk_season risk_season1
```

```
tab risk_season  
tab risk_season if _d==1  
tab cause_endpoint_enc risk_season if _d==1
```

```
*****PROTECTION PERIOD VARIABLE*****
```

```
stsplit treat_split, at(15795 16466 16526 16692 16914 17013 17236 17803 18020 18078 19018  
19097) after(capture_date==1/1/1960)
```

```
***Generating liberalized killing treatment binary variable (1 if lib kill period)
```

```
gen lib_kill = 0  
replace lib_kill = 1 if treat_split==15795 | treat_split==16526 | treat_split==16914 |  
treat_split==17236 | treat_split==18020 | treat_split==19018
```

```

order lib_kill treat_split, after(cause_endpoint_agg)

tab lib_kill if _d==1
/*346 deaths outside lib kill periods; 109 within*/
tab cause_endpoint2 lib_kill if _d==1
/*very few (<10) observations for collision (3) & unknown (7) and x-ltf (6) within lib kill
periods*/
tab risk_season lib_kill if _d==1

*censoring all other created spells
replace event=0 if event==.

***SAVE DATASET***
save "Supp_Dataset_expanded_xLTF.dta", replace

tab cause_endpoint_enc if _d==1

*****
*JOINT ST COX MODELS FOR ALL CAUSE-SPECIFIC HAZARDS*****
*****

*Checking stset by wolf_ID_exp (for multiple records)
stset date_endpoint, failure(event) exit(failure) origin(time capture_date) id(wolf_ID_exp)
stdes
stsum

*****
***CAUSE-SPECIFIC HAZARD RATES FOR ALL CAUSES SIMULTANEOUSLY*****
*Fit all in same model (basically same results as with separate PH models)
stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2

```

```

1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
1.risk_season#8.cause_endpoint2 2.risk_season#2.cause_endpoint2
2.risk_season#3.cause_endpoint2 2.risk_season#4.cause_endpoint2
2.risk_season#5.cause_endpoint2 2.risk_season#6.cause_endpoint2
2.risk_season#7.cause_endpoint2 2.risk_season#8.cause_endpoint2
3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2
3.risk_season#8.cause_endpoint2, efron strata(cause_endpoint2 lib_kill) robust cluster(wolf_ID)
estimates store fullperiod_KFstrat
estat ic

```

```

/*Sample code for checking assumptions*/

```

```

estat phtest, log detail
estat phtest, plot(1.risk_season#4.cause_endpoint2)
estat phtest, plot(2.risk_season#4.cause_endpoint2)
estat phtest, plot(1.risk_season#5.cause_endpoint2)
estat phtest, plot(2.risk_season#5.cause_endpoint2)
estat phtest, plot(3.risk_season#5.cause_endpoint2)
estat phtest, plot(2.risk_season#7.cause_endpoint2)

```

```

stphplot if cause_endpoint2==2, by(risk_season) nolnt
stphplot if cause_endpoint2==3, by(risk_season) nolnt
stphplot if cause_endpoint2==4, by(risk_season) nolnt
stphplot if cause_endpoint2==5, by(risk_season) nolnt
stphplot if cause_endpoint2==6, by(risk_season) nolnt
stphplot if cause_endpoint2==7, by(risk_season) nolnt
stphplot if cause_endpoint2==8, by(risk_season) nolnt

```

```

*WITH TVCs

```

```

stcox 1.risk_season#2.cause_endpoint2 1.risk_season#3.cause_endpoint2
1.risk_season#4.cause_endpoint2 1.risk_season#5.cause_endpoint2
1.risk_season#6.cause_endpoint2 1.risk_season#7.cause_endpoint2
1.risk_season#8.cause_endpoint2 2.risk_season#2.cause_endpoint2
2.risk_season#3.cause_endpoint2 2.risk_season#4.cause_endpoint2
2.risk_season#5.cause_endpoint2 2.risk_season#6.cause_endpoint2
2.risk_season#7.cause_endpoint2 2.risk_season#8.cause_endpoint2
3.risk_season#2.cause_endpoint2 3.risk_season#3.cause_endpoint2
3.risk_season#4.cause_endpoint2 3.risk_season#5.cause_endpoint2
3.risk_season#6.cause_endpoint2 3.risk_season#7.cause_endpoint2
3.risk_season#8.cause_endpoint2, tvc(1.risk_season#5.cause_endpoint2
1.risk_season#8.cause_endpoint2 2.risk_season#5.cause_endpoint2
2.risk_season#7.cause_endpoint2 3.risk_season#5.cause_endpoint2) texp(ln(_t)) efron
strata(cause_endpoint2 lib_kill) robust cluster(wolf_ID)
estimates store fullperiod_KFTVCstrat
estat ic

estout fullperiod_KFTVCstrat, eform cells("b(star fmt(3)) ci_l ci_u" se(par fmt(2))) starlevels(*
0.10 ** .05 *** 0.01) stats(ll aic bic, labels("Log likelihood" "AIC" "BIC")) legend

```